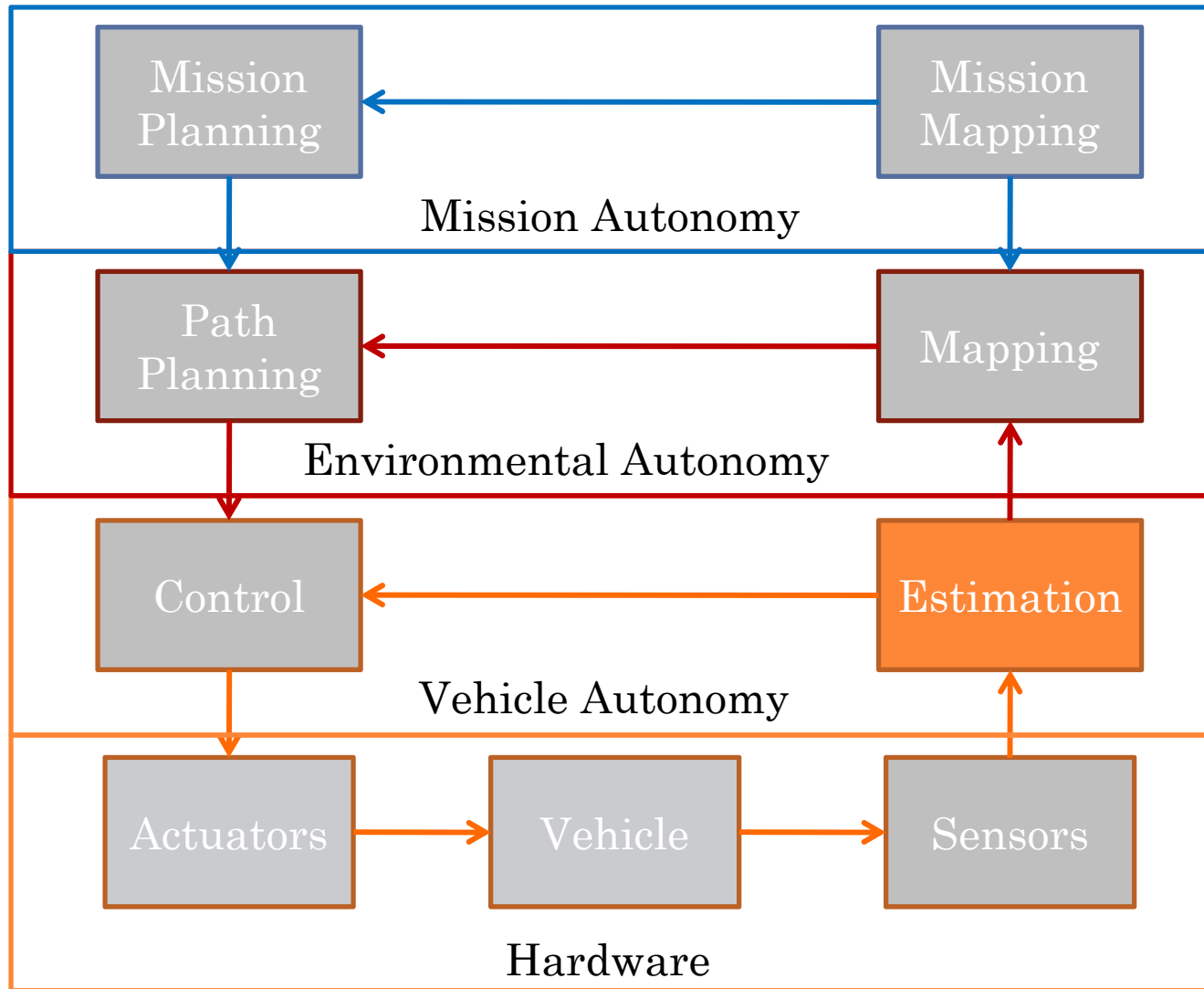


ME 597: AUTONOMOUS MOBILE ROBOTICS SECTION 7 – ESTIMATION I

Prof. Steven Waslander

COMPONENTS



OUTLINE

- Bayes Filter Framework
- Kalman Filter
- Extended Kalman Filter
- Particle Filter

BAYES FILTER

- The Bayes Filter forms the foundation for all other filters in this class
 - As described in background slides, Bayes rule is the right way to incorporate new probabilistic information into an existing, prior estimate
 - The resulting filter definition can be implemented directly for discrete state systems
 - For continuous states, need additional assumptions, additional structure to solve the update equations analytically

BAYES FILTER

- State x_t
 - All aspects of the vehicle and its environment that can impact the future
 - **Assume the state is complete**
- Control inputs u_t
 - All elements of the vehicle and its environment that can be controlled
- Measurements y_t
 - All elements of the vehicle and its environment that can be sensed
- Note: sticking with Thrun, Burgard, Fox notation
 - Discrete time index t
 - Initial state is x_0
 - First, apply control action u_1
 - Move to state x_1
 - Then, take measurement y_1

BAYES FILTER

- Motion Modeling

- Complete state:

- At each time t , x_{t-1} is a sufficient summary of all previous inputs and measurements

$$p(x_t | x_{0:t-1}, y_{1:t-1}, u_{1:t}) = p(x_t | x_{t-1}, u_t)$$

- Application of Conditional Independence
 - No additional information is to be had by considering previous inputs or measurements
 - Referred to as the Markov Assumption
 - Motion model is a Markov Chain

BAYES FILTER

- Measurement Modeling

- Complete state:
 - Current state is sufficient to model all previous states, measurements and inputs

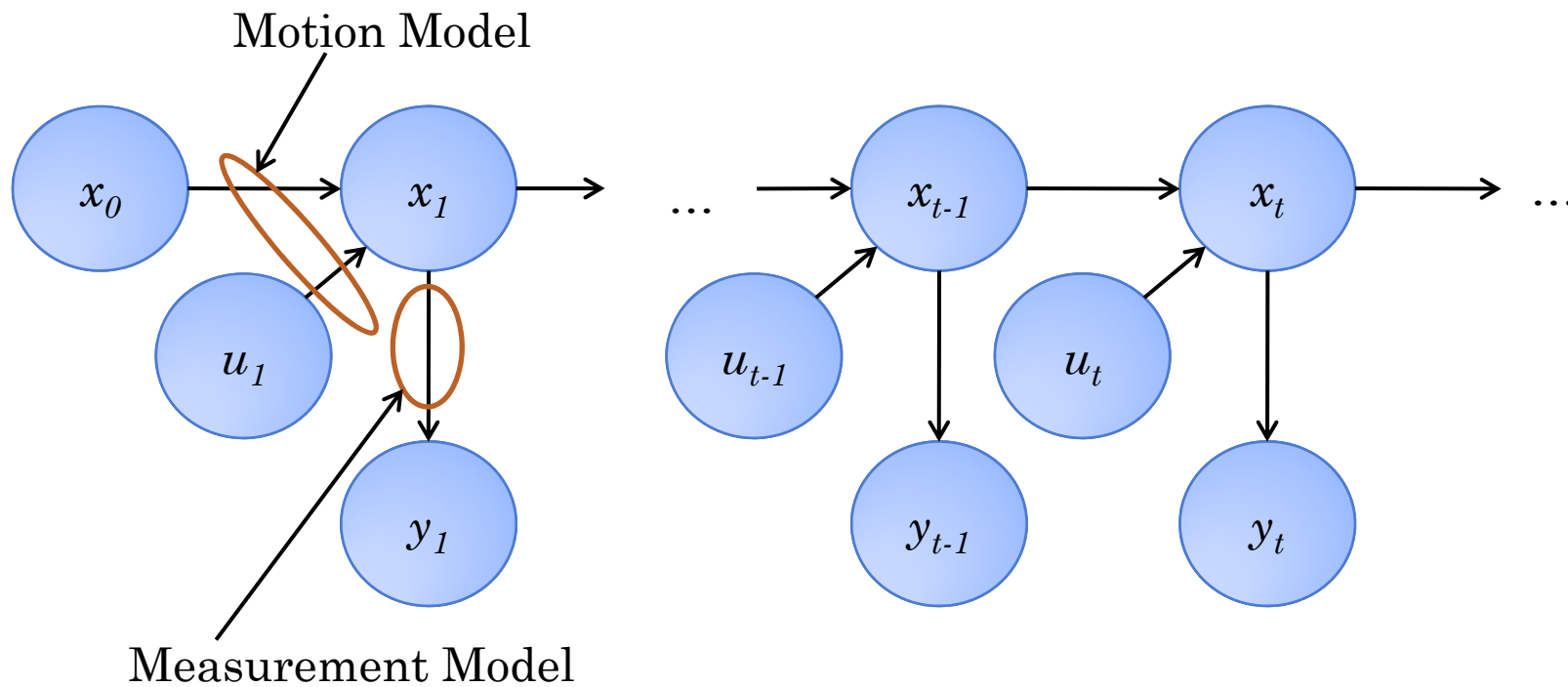
$$p(y_t | x_{0:t}, y_{1:t-1}, u_{1:t}) = p(y_t | x_t)$$

- Again, conditional independence
- Recall, in standard LTI state space model, measurement model may also depend on the current input

BAYES FILTER

- Combined Model

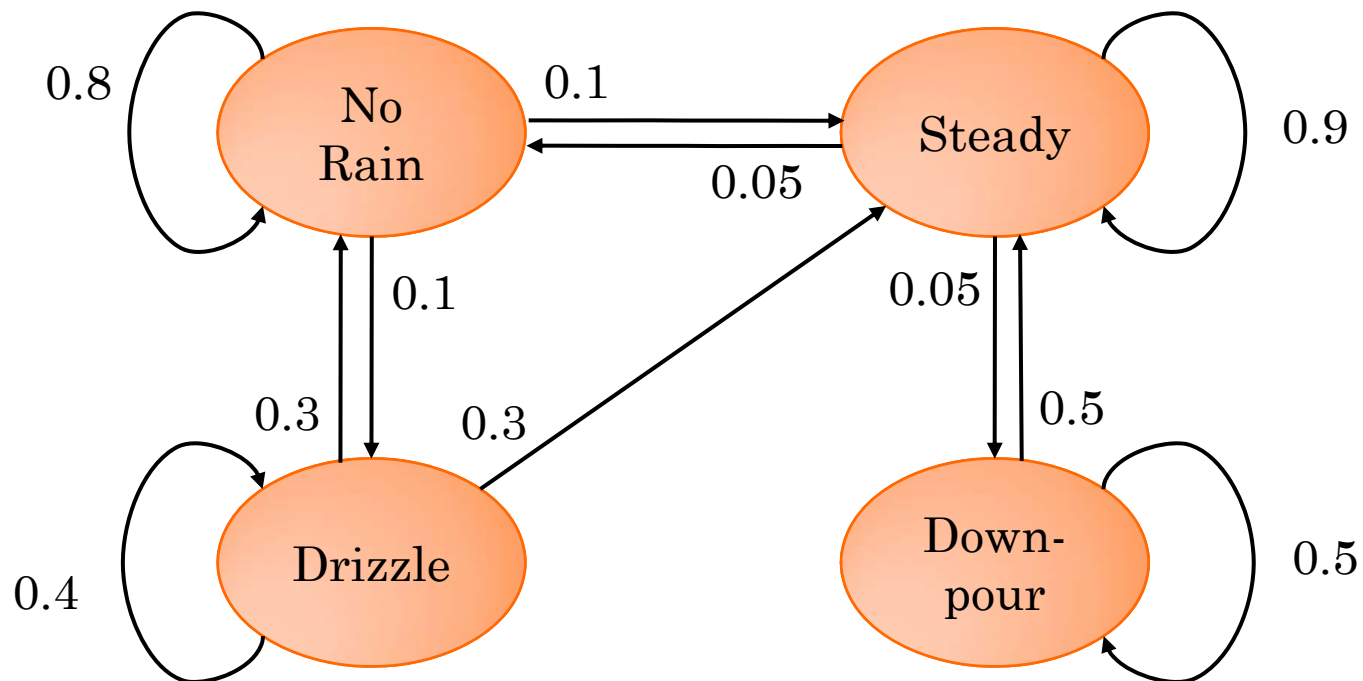
- Referred to as Hidden Markov Model (HMM) or Dynamic Bayes Network (DBN)



BAYES FILTER

- Example Discrete State Motion Model:

- States: $\{No\ Rain, Drizzle, Steady, Downpour\}$
- Inputs: None



BAYES FILTER

- For discrete states, the motion model can be written in matrix form
 - For each input u_t , the $n \times n$ motion model matrix is

$$p(x_t | u_t = u, x_{t-1}) = \begin{bmatrix} p(x_t = x_1 | x_{t-1} = x_1) & p(x_t = x_1 | x_{t-1} = x_2) & \cdots \\ p(x_t = x_2 | x_{t-1} = x_1) & p(x_t = x_2 | x_{t-1} = x_2) & \cdots \\ \vdots & \vdots & \ddots \end{bmatrix}$$

- Each row defines the probabilities of transitioning to state x_t from all possible states x_{t-1}
- Each column defines the probabilities of transitioning to any state x_t from a specific state x_{t-1}
- Again, the columns must sum to 1

BAYES FILTER

- Example:
 - Motion Model in Matrix Form
 - No inputs, one matrix

$$p(x_t | u_t, x_{t-1}) = \begin{matrix} & \overbrace{\hspace{10em}}^{x_{t-1}} & & \\ & \left[\begin{array}{cccc} 0.8 & 0.3 & 0.05 & 0 \\ 0.1 & 0.4 & 0 & 0 \\ 0.1 & 0.3 & 0.9 & 0.5 \\ 0 & 0 & 0.05 & 0.5 \end{array} \right] & \left. \vphantom{\begin{array}{cccc} 0.8 & 0.3 & 0.05 & 0 \\ 0.1 & 0.4 & 0 & 0 \\ 0.1 & 0.3 & 0.9 & 0.5 \\ 0 & 0 & 0.05 & 0.5 \end{array}} \right\} x_t \end{matrix}$$

BAYES FILTER

- Example Measurement Model:

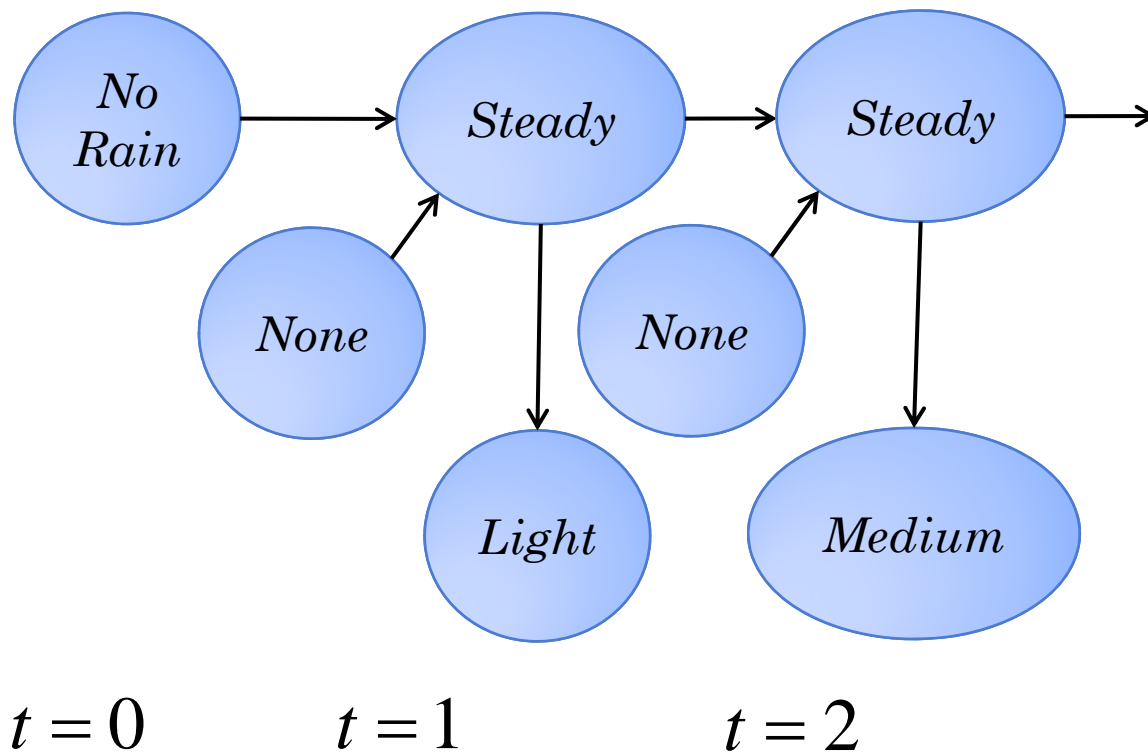
- States: $\{No\ Rain, Drizzle, Steady, Downpour\}$
- Measurements: $\{Dry, Light, Medium, Heavy\}$

$$p(y_t | x_t) = \begin{matrix} & \underbrace{\hspace{10em}}_{x_t} & & & & \\ & \left[\begin{array}{cccc} 0.95 & 0.1 & 0 & 0 \\ 0.05 & 0.8 & 0.15 & 0 \\ 0 & 0.1 & 0.7 & 0.1 \\ 0 & 0 & 0.15 & 0.9 \end{array} \right] & \left. \vphantom{\begin{array}{cccc} 0.95 & 0.1 & 0 & 0 \\ 0.05 & 0.8 & 0.15 & 0 \\ 0 & 0.1 & 0.7 & 0.1 \\ 0 & 0 & 0.15 & 0.9 \end{array}} \right\} & y_t \end{matrix}$$

- Again, the columns sum to 1

BAYES FILTER

- Example System Evolution



BAYES FILTER

○ Aim of Bayes Filter

- To estimate the current state of the system based on all known inputs and measurements.
- That is, to define a belief about the current state using all available information:

$$bel(x_t) = p(x_t | y_{1:t}, u_{1:t})$$

- Known as belief, state of knowledge, information state
- Depends on every bit of information that exists up to time t
- Can also define a belief prior to measurement y_t

$$\overline{bel}(x_t) = p(x_t | y_{1:t-1}, u_{1:t})$$

- Known as prediction, predicted state

BAYES FILTER

○ Problem Statement

- Given a prior for the system state

$$p(x_0)$$

- Given motion and measurement models

$$\overbrace{p(x_t | x_{t-1}, u_t)} \quad \overbrace{p(y_t | x_t)}$$

- Given a sequence of inputs and measurements

$$u_{1:t} = \{u_1, \dots, u_t\}, \quad y_{1:t} = \{y_1, \dots, y_t\}$$

- Estimate the current state distribution (form a belief about the current state)

$$bel(x_t) = p(x_t | y_{1:t}, u_{1:t})$$

BAYES FILTER

○ Bayes Filter Algorithm

- At each time step, t , for all possible values of the state x

1. Prediction update (Total probability)

$$\overline{bel}(x_t) = \int p(x_t | u_t, x_{t-1}) bel(x_{t-1}) dx_{t-1}$$

2. Measurement update (Bayes Theorem)

$$bel(x_t) = \eta p(y_t | x_t) \overline{bel}(x_t)$$

- η is a normalizing constant that does not depend on the state (will become apparent in derivation)
- Recursive estimation technique

BAYES FILTER

- Recall Bayes Theorem

$$p(a | b) = \frac{p(b | a)p(a)}{p(b)}$$

- Terminology

$$posterior = \frac{likelihood \cdot prior}{evidence}$$

BAYES FILTER

○ Derivation

- Proof by induction
 - Demonstrate that belief at time t can be found using belief at time $t-1$, input at t and measurement at t
- Initially

$$bel(x_0) = p(x_0)$$

- At time t , Bayes Theorem relates x_t, y_t

$$bel(x_t) = p(x_t | y_{1:t}, u_{1:t}) = p(x_t | y_t, y_{1:t-1}, u_{1:t})$$

$$= \frac{\overset{\textcircled{1}}{p(y_t | x_t, y_{1:t-1}, u_{1:t})} \overset{\textcircled{2}}{p(x_t | y_{1:t-1}, u_{1:t})}}{\underset{\textcircled{3}}{p(y_t | y_{1:t-1}, u_{1:t})}}$$

BAYES FILTER

○ Derivation

1. Measurement model simplifies first numerator term

$$p(y_t | x_t, y_{1:t-1}, u_{1:t}) = p(y_t | x_t)$$

2. Second numerator term is definition of belief prediction

$$\overline{bel}(x_t) = p(x_t | y_{1:t-1}, u_{1:t})$$

3. Denominator is independent of state, and so is constant for each time step. Define the normalizer,

$$\eta = p(y_t | y_{1:t-1}, u_{1:t})^{-1}$$

BAYES FILTER

○ Derivation

- Summarizing the three substitutions

$$bel(x_t) = \eta p(y_t | x_t) \overline{bel}(x_t)$$

- This is exactly the measurement update step
- Requires the measurement y_t to be known
- However, we now need to find the belief prediction
 - Done using total probability, over previous state

$$\begin{aligned} \overline{bel}(x_t) &= p(x_t | y_{1:t-1}, u_{1:t}) \\ &= \int p(x_t | x_{t-1}, y_{1:t-1}, u_{1:t}) p(x_{t-1} | y_{1:t-1}, u_{1:t}) dx_{t-1} \end{aligned}$$

(1) (2)

BAYES FILTER

○ Derivation

1. This time, the motion model can be incorporated

$$p(x_t | x_{t-1}, y_{1:t-1}, u_{1:t}) = p(x_t | x_{t-1}, u_t)$$

2. And we note that the control input at time t does not affect the state at time $t-1$

$$\begin{aligned} p(x_{t-1} | y_{1:t-1}, u_{1:t}) &= p(x_{t-1} | y_{1:t-1}, u_{1:t-1}) \\ &= \text{bel}(x_{t-1}) \end{aligned}$$

BAYES FILTER

○ Derivation

- And so the prediction update is defined

$$\overline{bel}(x_t) = \int p(x_t | x_{t-1}, u_t) bel(x_{t-1}) dx_{t-1}$$

- Which completes the proof by induction
 - For this step, we need the control input to define the correct motion model distribution
- If state, measurements, inputs are discrete, can directly implement Bayes Filter
 - Prediction update is summation over discrete states
 - Measurement update is multiplication of two vectors

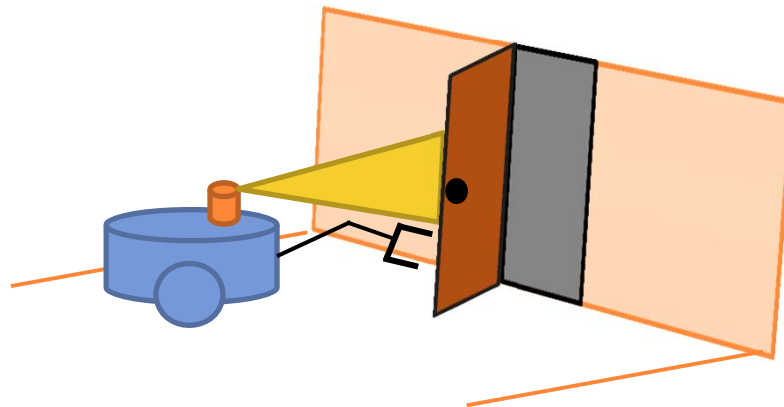
BAYES FILTER

- If state, measurement, inputs are continuous, must define model or approximation to enable computation
 - Kalman Filter:
 - Linear motion models
 - Linear measurement models
 - Additive Gaussian disturbance and noise distributions
 - Extended Kalman Filter/Unscented Kalman Filter:
 - Nonlinear motion models
 - Nonlinear measurement models
 - Additive Gaussian disturbance and noise distributions
 - Particle Filter:
 - (Dis)continuous motion models
 - (Dis)continuous measurement models
 - General disturbance and noise distributions

BAYES FILTER

○ Discrete Bayes Filter Example

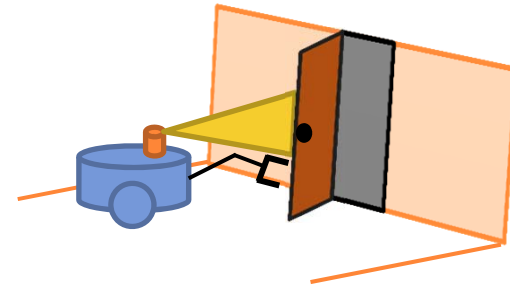
- Problem: Detect if a door is open/closed with a robot that can sense the door position and pull the door open



- State: $door = \{open, closed\}$
- State Prior (uniform):

$$p(x_0) = \begin{cases} p(open) = 0.5 \\ p(closed) = 0.5 \end{cases}$$

BAYES FILTER



○ Example

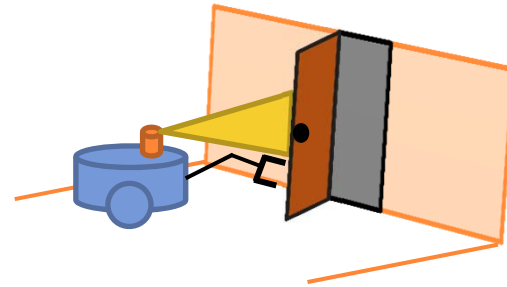
- Inputs: $arm_command = \{none, pull\}$
- Motion Model
 - If input = *none*, do nothing:

$$p(x_t | u_t = none, x_{t-1}) \rightarrow \begin{cases} p(open_t | none, open_{t-1}) = 1 \\ p(closed_t | none, open_{t-1}) = 0 \\ p(open_t | none, closed_{t-1}) = 0 \\ p(closed_t | none, closed_{t-1}) = 1 \end{cases}$$

- If input = *pull*, pull the door open

$$p(x_t | u_t = pull, x_{t-1}) \rightarrow \begin{cases} p(open_t | pull, open_{t-1}) = 1 \\ p(closed_t | pull, open_{t-1}) = 0 \\ p(open_t | pull, closed_{t-1}) = 0.8 \\ p(closed_t | pull, closed_{t-1}) = 0.2 \end{cases}$$

BAYES FILTER

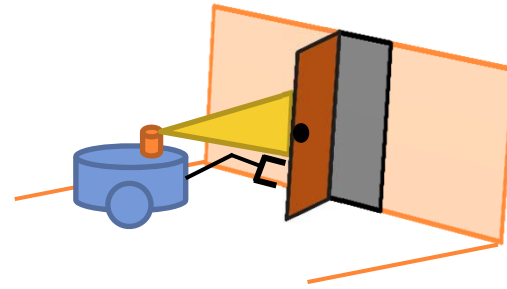


○ Example

- Measurements: $meas = \{sense_open, sense_closed\}$
- Measurement model (noisy door sensor):

$$p(y | x) \rightarrow \begin{cases} p(sense_open | open) = 0.6 \\ p(sense_open | closed) = 0.2 \\ p(sense_closed | open) = 0.4 \\ p(sense_closed | closed) = 0.8 \end{cases}$$

BAYES FILTER



○ Example

- At time step 1, input = *none*
- Perform state prediction update

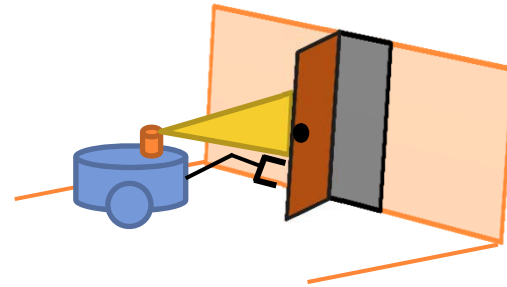
$$\begin{aligned}\overline{bel}(x_1) &= \int p(x_1 | u_0, x_0) bel(x_0) dx_0 \\ &= \sum p(x_1 | u_0, x_0) p(x_0)\end{aligned}$$

- Calculate belief prediction for each possible value of state

$$\begin{aligned}\overline{bel}(open_1) &= p(open_1 | none_1, open_0) bel(open_0) \\ &\quad + p(open_1 | none_1, closed_0) bel(closed_0) \\ &= 1 * 0.5 + 0 * 0.5 = 0.5\end{aligned}$$

$$\begin{aligned}\overline{bel}(closed_1) &= p(closed_1 | none_1, open_0) bel(open_0) \\ &\quad + p(closed_1 | none_1, closed_0) bel(closed_0) \\ &= 0 * 0.5 + 1 * 0.5 = 0.5\end{aligned}$$

BAYES FILTER



○ Example

- At time step 1, measurement $y_1 = \text{sense_open}$
- Measurement update

$$\text{bel}(x_1) = \eta p(y_1 | x_1) \overline{\text{bel}}(x_1)$$

- Calculate for each possible value of state

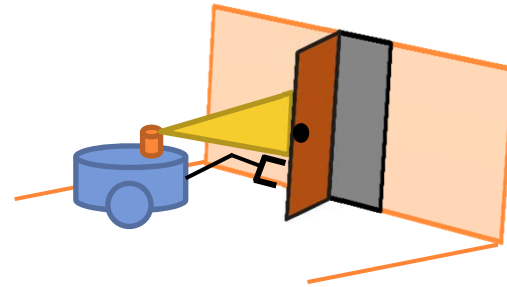
$$\begin{aligned} \text{bel}(\text{open}_1) &= \eta p(\text{sense_open}_1 | \text{open}_1) \overline{\text{bel}}(\text{open}_1) \\ &= \eta 0.6 \cdot 0.5 = 0.3\eta \end{aligned}$$

$$\begin{aligned} \text{bel}(\text{closed}_1) &= \eta p(\text{sense_open}_1 | \text{closed}_1) \overline{\text{bel}}(\text{closed}_1) \\ &= \eta 0.2 \cdot 0.5 = 0.1\eta \end{aligned}$$

- Calculate normalizer and solve for posterior

$$\eta = \frac{1}{0.3 + 0.1} = 2.5 \quad \longrightarrow \quad \begin{aligned} \text{bel}(\text{open}_1) &= 0.75 \\ \text{bel}(\text{closed}_1) &= 0.25 \end{aligned}$$

BAYES FILTER



○ Example

- At time step 2, a *pull* and a *sense_open*
- Then state propagation

$$\overline{bel}(open_2) = 1 \cdot 0.75 + 0.8 \cdot 0.25 = 0.95$$

$$\overline{bel}(closed_2) = 0 \cdot 0.75 + 0.2 \cdot 0.25 = 0.05$$

- And measurement update

$$bel(open_2) = \eta 0.6 \cdot 0.95 = 0.983$$

$$bel(closed_2) = \eta 0.2 \cdot 0.05 = 0.017$$

- In summary:
 - Uniform prior, do nothing, measure open: $bel(open_1) = 0.75$
 - Pull open, measure open: $bel(open_2) = 0.983$

BAYES FILTER

○ Example 2: Histogram Filter

- Motion of robot in a $n \times n$ grid

- State:

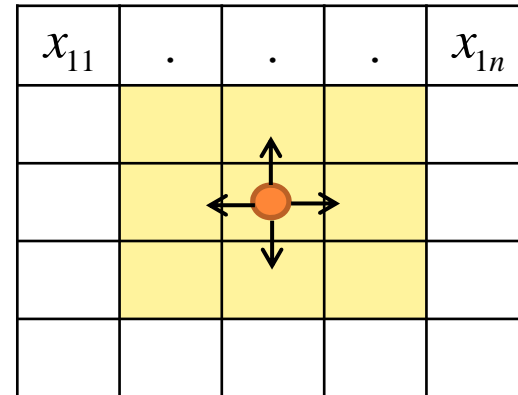
- Position = $\{x_{11}, x_{12}, \dots, x_{1n}, \dots, x_{nn}\}$

- Input:

- Move = {Up, Right, Down, Left}
- 40% chance the move does not happen
- Cannot pass through outer walls

- Measurement: Accurate to within 3X3 grid

$$p(y(i-1:i+1, j-1:j+1) | x(i, j)) = \begin{bmatrix} .11 & .11 & .11 \\ .11 & .12 & .11 \\ .11 & .11 & .11 \end{bmatrix}$$



BAYES FILTER

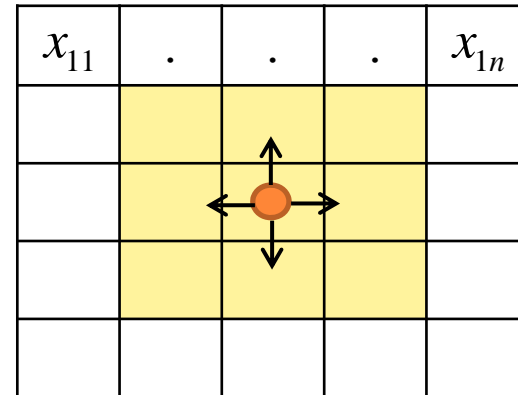
○ Example 2:

- Prior over states
 - Assume no information, uniform
 - Vector of length n^2

$$p(x_0) = \frac{1}{n^2} \mathbf{1}$$

- Motion model
 - Given a particular input and previous state, probability of moving to any other state
 - $n \times n$ state, one for each grid point
 - 4 input choices

$$p(x_t | x_{t-1}, u_t) \in [0, 1]^{n^2 \times n^2 \times 4}$$



BAYES FILTER

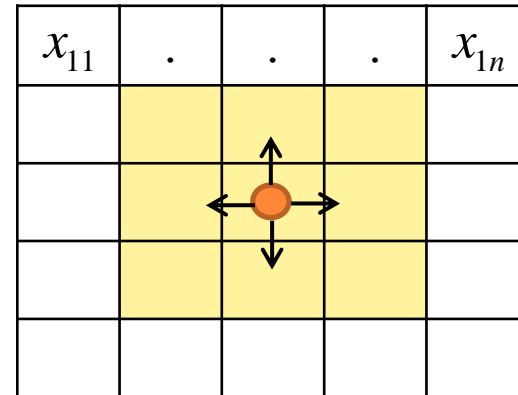
○ Example 2

- Measurement Model

- Given any current state, probability of a measurement
- Same number of measurements as states

$$p(y_t | x_t) \in [0,1]^{n^2 \times n^2}$$

- Same 3X3 matrix governs all interior points
- Boundaries cut off invalid measurements and require normalization
- Very simplistic and bloated model
 - Could replace with 2 separate states and measurements to perpendicular walls

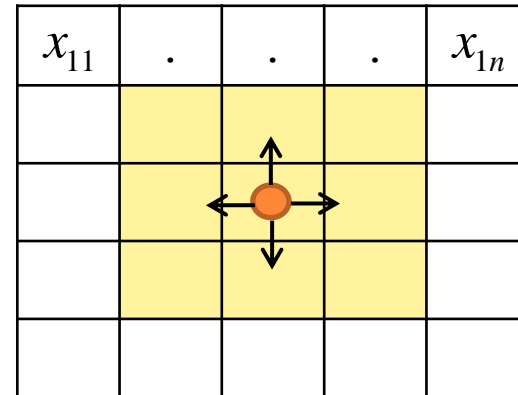


BAYES FILTER

○ Example 2 – Motion Model code

```
mot_mod = zeros(N,N,4);
for i=1:n
    for j=1:n
        cur = i+(j-1)*n;
        % Move up
        if (j > 1)
            mot_mod(cur-n,cur,1) = 0.6;
            mot_mod(cur,cur,1) = 0.4;
        else
            mot_mod(cur,cur,1) = 1;
        end
        % Move right
        if (i < n)
            mot_mod(cur+1,cur,2) = 0.6;
            mot_mod(cur,cur,2) = 0.4;
        else
            mot_mod(cur,cur,2) = 1;
        end
    end
end
```

...

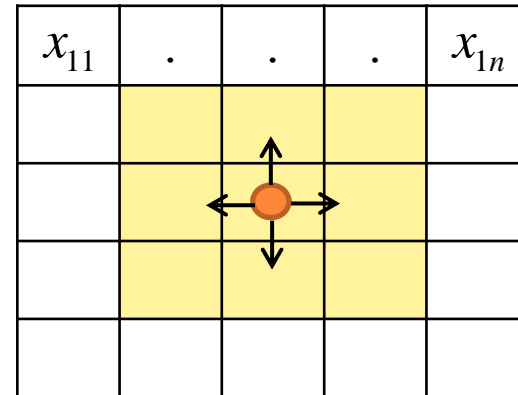


BAYES FILTER

○ Example 2 – Measurement Model

```
%% Create the measurement model
meas_mod_rel = [0.11 0.11 0.11;
                0.11 0.12 0.11;
                0.11 0.11 0.11];

% Convert to full measurement model
% p(y_t | x_t)
meas_mod = zeros(N,N);
% Fill in non-boundary measurements
for i=2:n-1
    for j=2:n-1
        cur = i+(j-1)*n;
        meas_mod(cur-n+[-1:1:1],cur) = meas_mod_rel(1,:);
        meas_mod(cur+[-1:1:1],cur) = meas_mod_rel(2,:);
        meas_mod(cur+n+[-1:1:1],cur) = meas_mod_rel(3,:);
    end
end
...
```

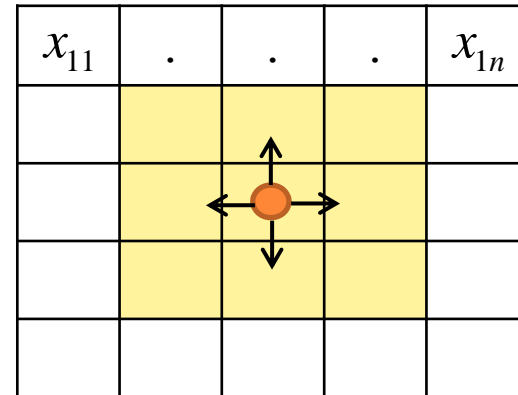


BAYES FILTER

- Example 2 – Makin' movies!

```
videoobj=VideoWriter('bayesgrid.mp4','MPEG-4');  
truefps = 1;  
videoobj.FrameRate = 10; %Anything less than 10 fps fails.  
open(videoobj);
```

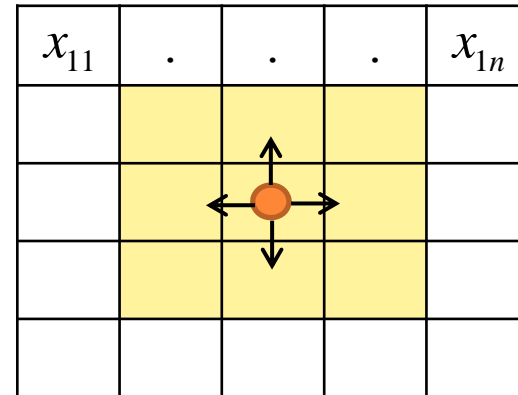
```
figure(1);clf; hold on;  
beliefs = reshape(bel,n,n);  
imagesc(beliefs);  
plot(pos(2),pos(1),'ro','MarkerSize',6,'LineWidth',2)  
colormap(bone);  
title('True state and beliefs')  
F = getframe;  
% Dumb hack to get desired framerate  
for dumb=1:floor(10/truefps)  
    writeVideo(videoobj, F);  
end  
...
```



BAYES FILTER

○ Example 2 – Simulation code

```
%Main Loop
for t=1:T
    %% Simulation
    % Select motion input
    u(t) = ceil(4*rand(1));
    % Select a motion
    thresh = rand(1);
    new_x = find(cumsum(squeeze(mot_mod(:, :, u(t))) * x(:, t)) > thresh, 1);
    % Move vehicle
    x(new_x, t+1) = 1;
    % Take measurement
    thresh = rand(1);
    new_y = find(cumsum(meas_mod(:, :) * x(:, t+1)) > thresh, 1);
    y(new_y, t) = 1;
    % Store for plotting
    ...
end
```



BAYES FILTER

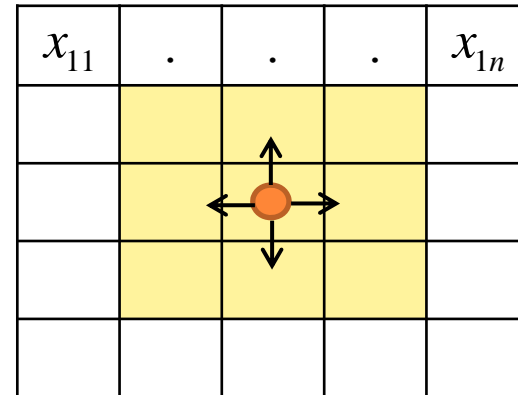
○ Example 2 – Bayes Filter

```
...
%% Bayesian Estimation
% Prediction update
belp = squeeze(mot_mod(:, :, u(t))) * bel;

% Measurement update
bel = meas_mod(new_y, :)' . * belp;
bel = bel / norm(bel);

[pmax y_bel(t)] = max(bel);

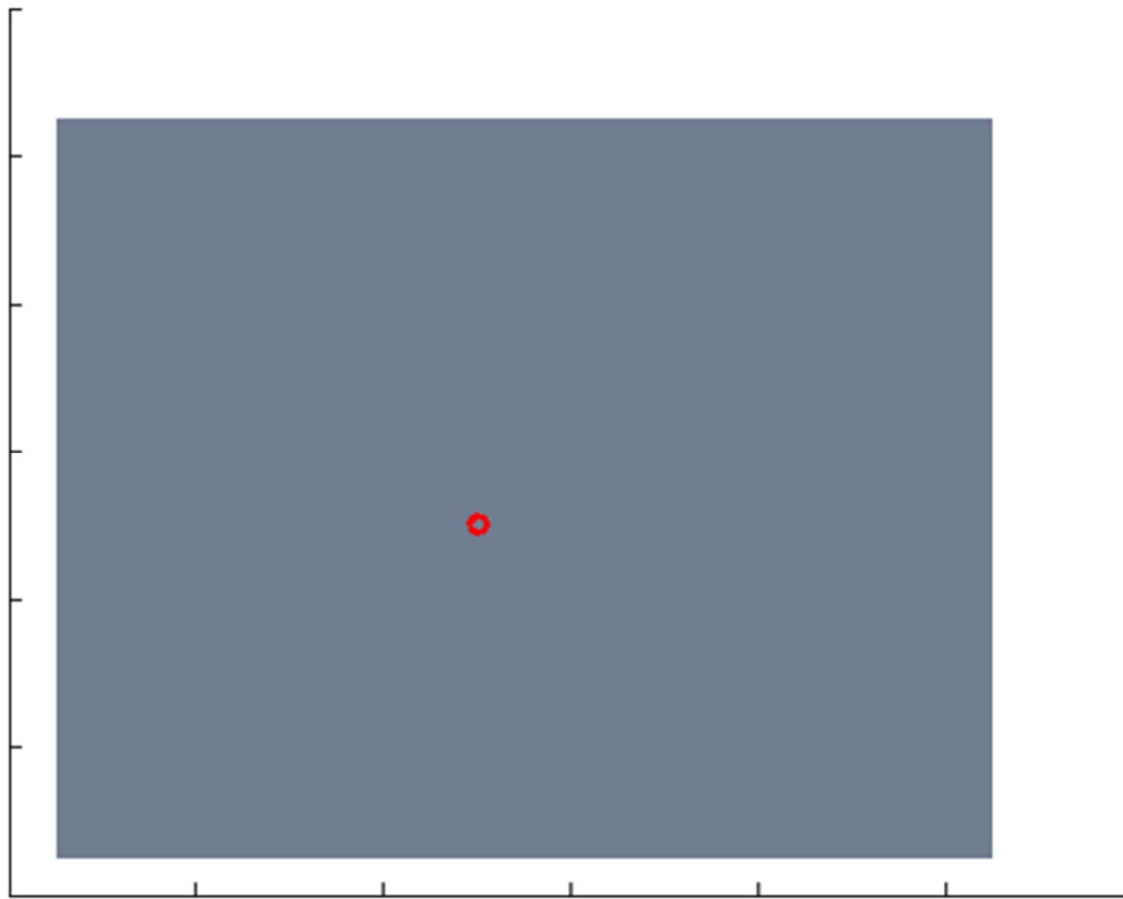
%% Plot beliefs
...
```



BAYES FILTER

- Example 2:

- Results

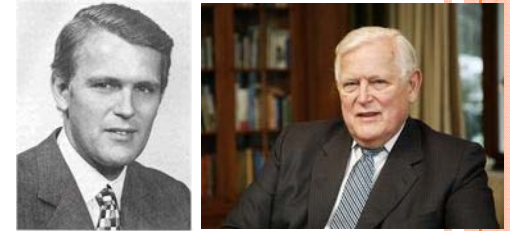


OUTLINE

- Bayes Filter Framework
- Kalman Filter
- Extended Kalman Filter
- Particle Filter

KALMAN FILTER

- Rudolf Kalman 1960 (BS, MS: MIT, PhD: Columbia)



The filter is named after [Rudolf E. Kalman](#), though [Thorvald Nicolai Thiele^{\[1\]}](#) and [Peter Swerling](#) developed a similar algorithm earlier. [Stanley F. Schmidt](#) is generally credited with developing the first implementation of a Kalman filter. It was during a visit of Kalman to the [NASA Ames Research Center](#) that he saw the applicability of his ideas to the problem of trajectory estimation for the [Apollo program](#), leading to its incorporation in the Apollo navigation computer. The filter was developed in papers by Swerling (1958), Kalman (1960), and Kalman and Bucy (1961).

Wikipedia

- Discrete version (Kalman filter)
- Continuous version (Kalman-Bucy filter)
- Many other versions, improvements, modifications

KALMAN FILTER

- Kalman Filter Modeling Assumption

- Continuous state, inputs, measurements
- Prior over the state is Gaussian

$$p(x_0) \sim N(\mu_0, \Sigma_0)$$

- Motion model, linear with additive Gaussian disturbances

$$x_t = A_t x_{t-1} + B_t u_t + \varepsilon_t \quad \varepsilon_t \sim N(0, R_t)$$

- Often, robotics systems are more easily described in continuous domain
 - Convert to discrete time using matrix exponential
 - Matlab contains tools to perform this conversion (c2d, d2c)

KALMAN FILTER

- Kalman Filter Modeling Assumption

- Measurement model also linear with additive Gaussian noise

$$y_t = C_t x_t + \delta_t \quad \delta_t \sim N(0, Q_t)$$

- Can add in input dependence to match up with controls literature

$$y_t = C_t x_t + D_t u_t + \delta_t$$

KALMAN FILTERING

○ Full Model

- State prior

$$p(x_0) \sim N(\mu_0, \Sigma_0)$$

- Motion model

$$x_t = A_t x_{t-1} + B_t u_t + \varepsilon_t \quad \varepsilon_t \sim N(0, R_t)$$

- Measurement model

$$y_t = C_t x_t + \delta_t \quad \delta_t \sim N(0, Q_t)$$

KALMAN FILTER

- Assume belief is Gaussian at time t

$$bel(x_t) \sim N(\mu_t, \Sigma_t)$$

- μ_t is the best estimate of the current state at time t
- Σ_t is the covariance, indicating the certainty in the current estimate

- Will be able to demonstrate the predicted belief at the next time step is Gaussian

$$\overline{bel}(x_{t+1}) \sim N(\overline{\mu}_{t+1}, \overline{\Sigma}_{t+1})$$

- And that the belief at next time step is also Gaussian

$$bel(x_{t+1}) \sim N(\mu_{t+1}, \Sigma_{t+1})$$

KALMAN FILTER

- Goal:

- To find belief over state as accurately as possible given all available information
 - Minimize the mean square error of the estimate (MMSE estimator)

$$\min E[(\mu_t - x_t)^2]$$

- Same as least square problem
- Using an unbiased estimator

$$E[\mu_t - x_t] = 0$$

- On average, your estimate is correct!

KALMAN FILTER

○ Goal

- The MMSE estimate can be written as

$$\min E[(x_t - \mu_t)^T (x_t - \mu_t)]$$

- And is equivalent to minimizing the trace of the error covariance matrix

$$\min \text{tr}(\Sigma_t)$$

- Proof:

$$\begin{aligned} E[(x_t - \mu_t)^T (x_t - \mu_t)] &= E\left[\sum_i (x_{t,i} - \mu_{t,i})^2\right] \\ &= E[\text{tr}((x_t - \mu_t)(x_t - \mu_t)^T)] \\ &= \text{tr}(E[(x_t - \mu_t)(x_t - \mu_t)^T]) \\ &= \text{tr}(\Sigma_t) \end{aligned}$$

KALMAN FILTER

○ Kalman Filter Algorithm

- At each time step, t , update both sets of beliefs
 1. Prediction update

$$\bar{\mu}_t = A_t \mu_{t-1} + B_t u_t$$

$$\bar{\Sigma}_t = A_t \Sigma_{t-1} A_t^T + R_t$$

2. Measurement update

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

$$\mu_t = \bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t)$$

$$\Sigma_t = (I - K_t C_t) \bar{\Sigma}_t$$

- Kalman Gain, K_t
 - Blending factor between prediction and measurement

KALMAN FILTER

○ Example

- Temperature control
 - State is current temperature difference with outside
 - One dimensional example
 - Prior: fairly certain of current temperature difference

$$\mu_0 = 10$$

$$\Sigma_0 = 1$$

- Motion Model: Decaying temperature + furnace input + disturbances (opening doors, outside effects)

$$dt = 0.1$$

$$x_t = 0.8x_{t-1} + 3u_t + r_t$$

$$A = 0.8, \quad B = 3$$

$$r_t \sim N(0, 2)$$

KALMAN FILTER

○ Example

- Measurement Model

- Directly measure the current temperature difference

$$y(t) = x(t) + \delta_t$$

$$\delta_t \sim N(0,4)$$

- Controller design

- Bang bang control, based on current estimate of temperature difference

$$u(t) = \begin{cases} 1 & \mu_t < 2 \\ 0 & \mu_t > 10 \\ u(t-1) & \text{otherwise} \end{cases}$$

KALMAN FILTER

○ Example

• Simulation

```
for t=1:length(T)
% Select control action
    if (t>1) u(t)=u(t-1); end
    if (mu > 10)
        u(t) = 0;
    elseif (mu < 2);
        u(t) = 1;
    end

% Update state
    e = sqrt(R)*randn(1);
    x(t+1) = A*x(t)+ B*u(t) + e;

% Determine measurement
    d = sqrt(Q)*randn(1);
    y(t) = C*x(t+1) + d;
```

KALMAN FILTER

○ Example

- Estimation

```
% Prediction update
mup = A*mu + B*u(t);
Sp = A*S*A' + R;

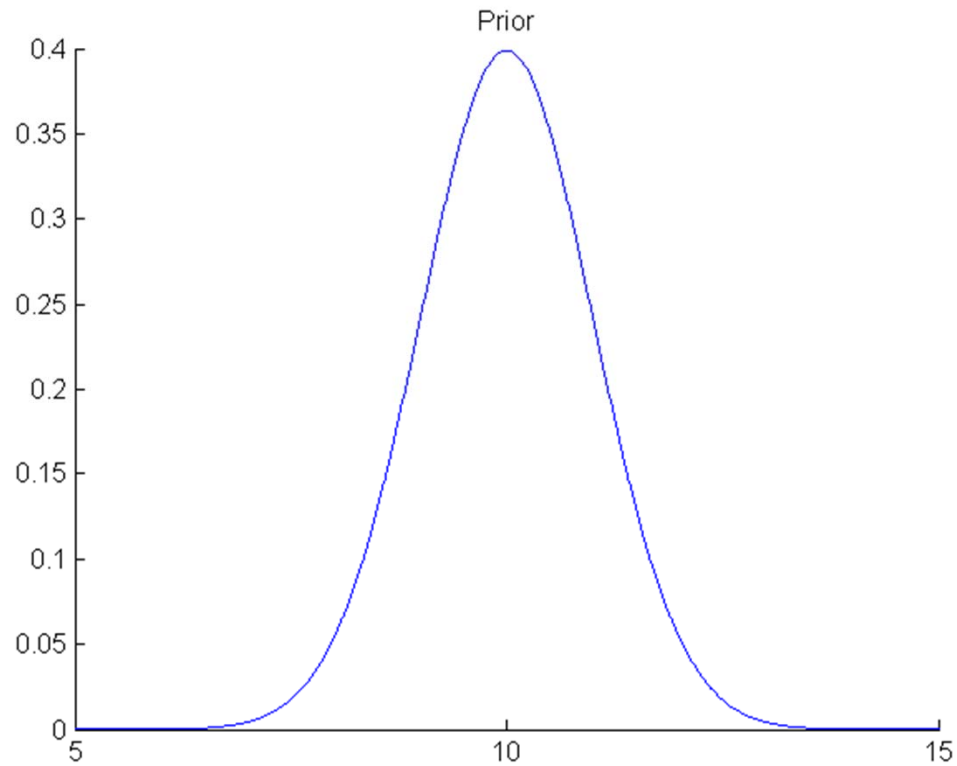
% Measurement update
K = Sp*C'*inv(C*Sp*C'+Q);
mu = mup + K*(y(t)-C*mup);
S = (1-K*C)*Sp;
```

- Matrix inverse $O(n^2.4)$, matrix multiplication $O(n^2)$
- When implementing in Matlab, `inv()` performs matrix inverse for you
- For embedded code, many libraries exist
 - Try Gnu Scientific Library, easy starting point

KALMAN FILTER

○ Example

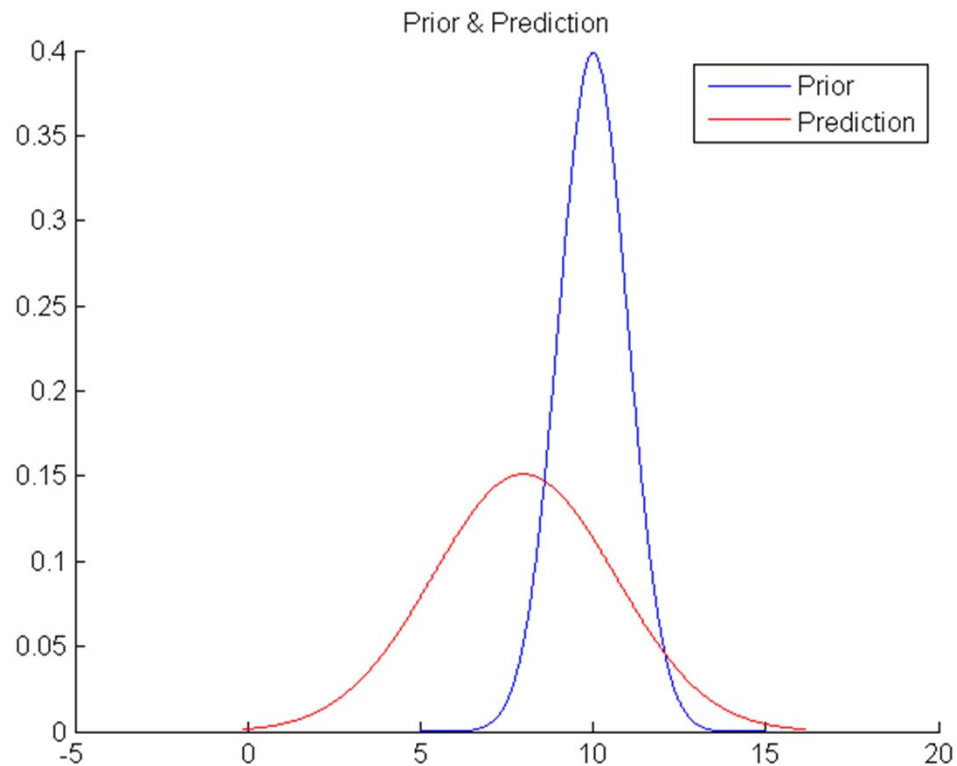
- Beliefs during the first time step
 - Prior



KALMAN FILTER

○ Example

- Beliefs during the first time step
 - Prediction Update: increased variance, shifted mean



$$\bar{\mu}_t = A_t \mu_{t-1} + B_t u_t$$

$$\bar{\Sigma}_t = A_t \Sigma_{t-1} A_t^T + R_t$$

$$x_t = 0.8x_{t-1} + 3u_t + r_t$$

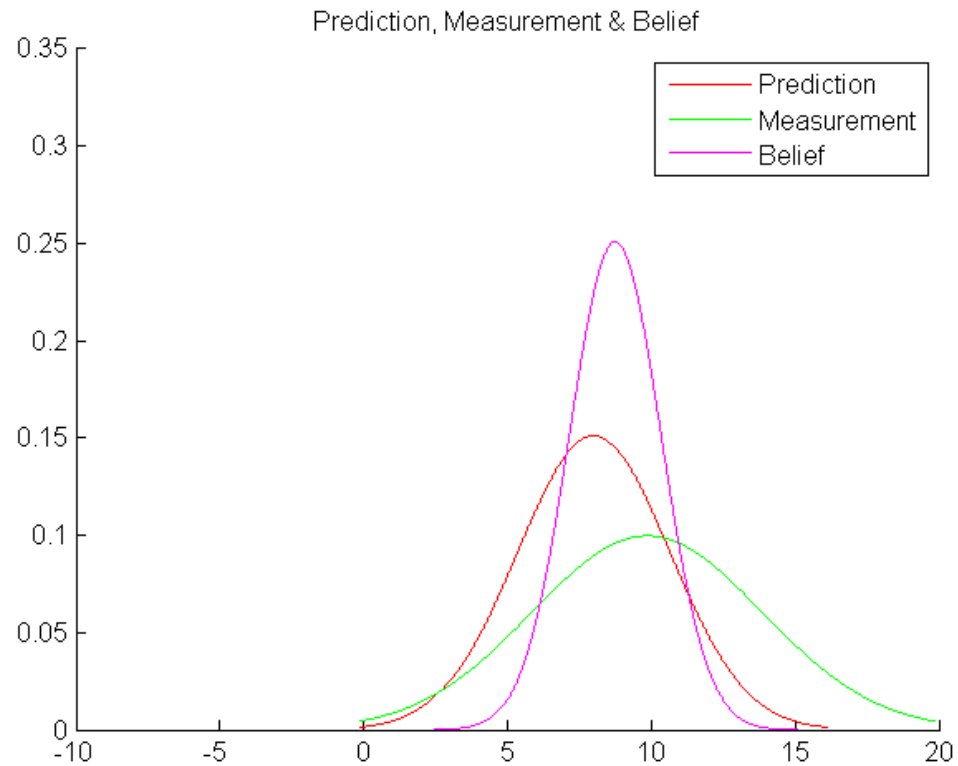
$$A = 0.8, \quad B = 3$$

$$r_t \sim N(0, 2)$$

KALMAN FILTER

○ Example

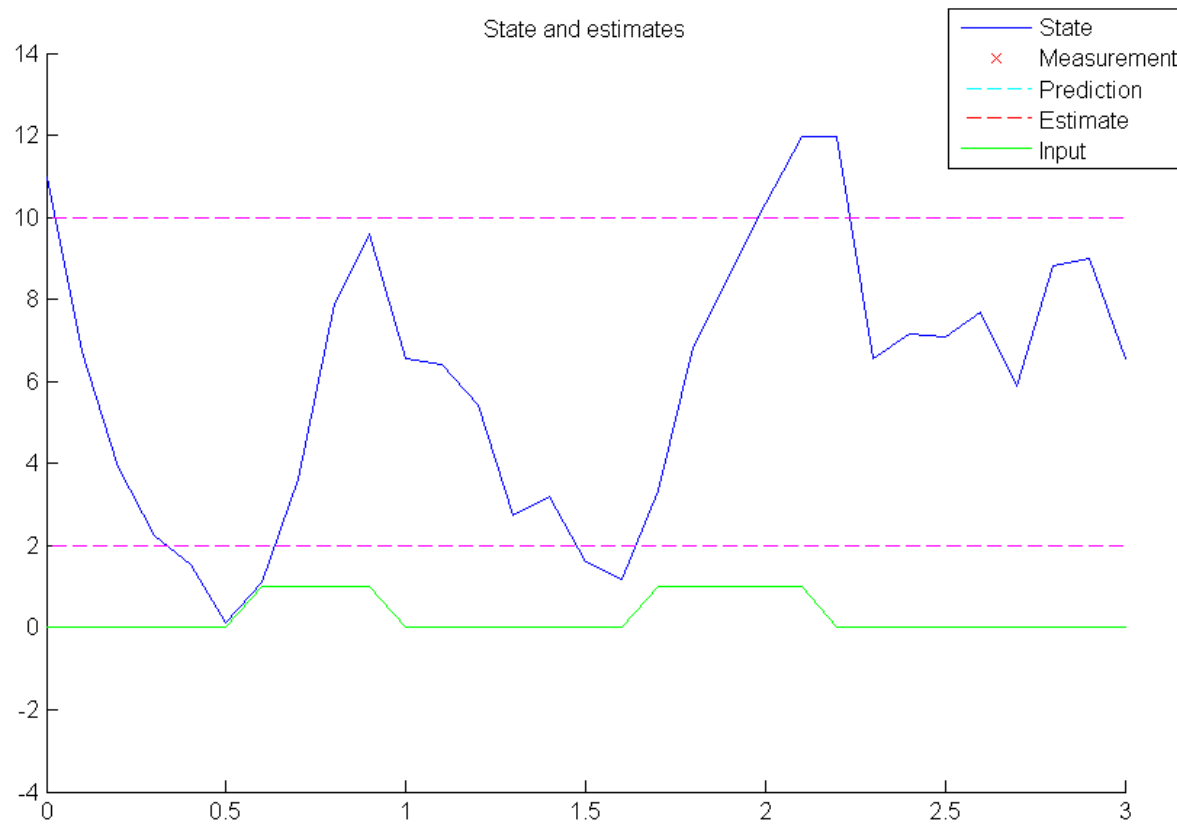
- Beliefs after the first time step
 - Measurement update: decreased variance



KALMAN FILTER

○ Example

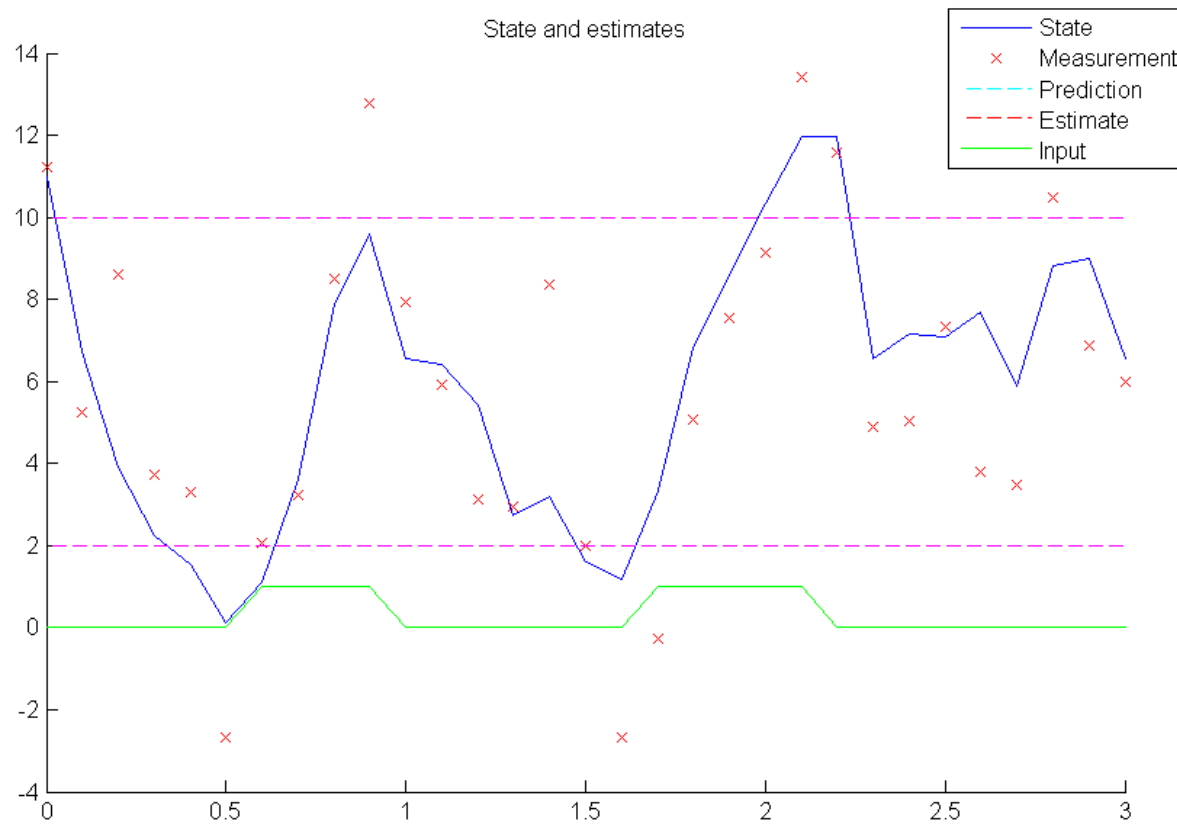
- Trajectories over time



KALMAN FILTER

○ Example

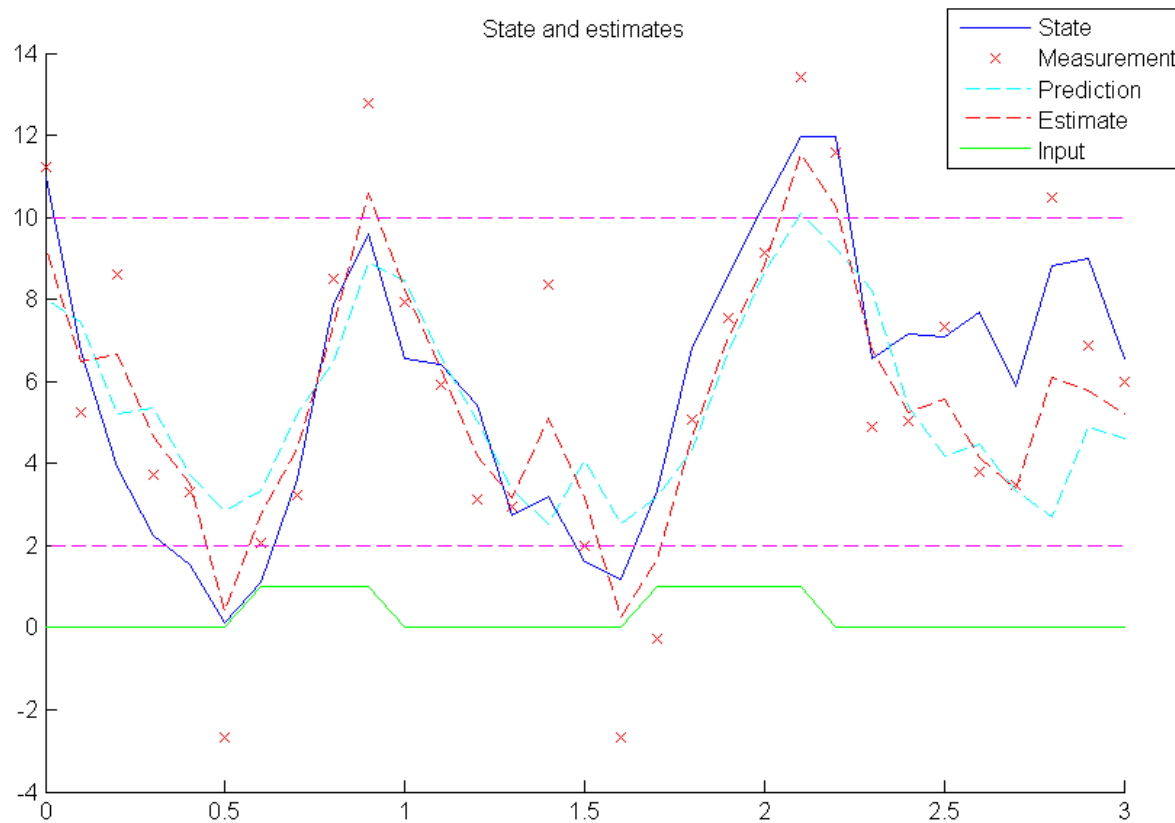
- Measurements over time



KALMAN FILTER

○ Example

- Estimates over time



KALMAN FILTER

- Recall Goal:

- To find belief over state as accurately as possible given all available information
 - Minimize the mean square error of the estimate (MMSE estimator)

$$\min E[(\mu_t - x_t)^2]$$

- Same as least square problem
- Using an unbiased estimator

$$E[\mu_t - x_t] = 0$$

- On average, your estimate is right!

KALMAN FILTER

○ Derivation

- Define the innovation
 - The difference between the measurement and the expected measurement given the predicted state and the measurement model

$$\text{Innovation}_t = y_t - C_t \bar{\mu}_t$$

- Assume the form of the estimator is a linear combination of the predicted belief and the innovation
 - The following form turns out to be unbiased

$$\mu_t = \bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t)$$

KALMAN FILTER

○ Steps

- Prediction update
 - Find update rule for mean, covariance of predicted belief, given input and motion model
- Measurement update
 - Solve MMSE optimization problem to find update rule for mean, covariance of belief given measurement model and measurement

KALMAN FILTER

○ Prediction Update

- Only new information is input u_t
- Prediction update is a linear transformation of belief at previous time step
 - Motion model is

$$x_t = A_t x_{t-1} + B_t u_t + \varepsilon_t$$

- Motion noise, previous belief are Gaussian so this is an addition of Gaussian distributions

$$bel(x_{t-1}) \sim N(\mu_{t-1}, \Sigma_{t-1}) \quad \varepsilon_t \sim N(0, R_t)$$

- Therefore the predicted mean and covariance are

$$\bar{\mu}_t = A_t \mu_{t-1} + B_t u_t + 0$$

$$\bar{\Sigma}_t = A_t \Sigma_{t-1} A_t^T + R_t$$

KALMAN FILTER

○ Measurement update

- First, let's define the form of the error covariance, substituting in the form of the mean update and the measurement model

$$\Sigma_t = E[(x_t - \mu_t)(x_t - \mu_t)^T]$$

- But, by the assumption of the form of the estimator, and the measurement model

$$\begin{aligned}x_t - \mu_t &= x_t - \bar{\mu}_t - K_t(y_t - C_t\bar{\mu}_t) \\ &= x_t - \bar{\mu}_t - K_t(C_t x_t + \delta_t - C_t\bar{\mu}_t) \\ &= (I - K_t C_t)(x_t - \bar{\mu}_t) - K_t \delta_t\end{aligned}$$

KALMAN FILTER

- Measurement update

- Next, reorganize terms of the covariance

$$\begin{aligned}\Sigma_t &= E[((I - K_t C_t)(x_t - \bar{\mu}_t) - K_t \delta_t) ((I - K_t C_t)(x_t - \bar{\mu}_t) - K_t \delta_t)^T] \\ &= E[((I - K_t C_t)(x_t - \bar{\mu}_t)) ((I - K_t C_t)(x_t - \bar{\mu}_t))^T] \\ &\quad - 2E[((I - K_t C_t)(x_t - \bar{\mu}_t)) K_t \delta_t] \\ &\quad + E[(K_t \delta_t) (K_t \delta_t)^T]\end{aligned}$$

- But the middle term is zero in expectation

$$\begin{aligned}\Sigma_t &= E[((I - K_t C_t)(x_t - \bar{\mu}_t)) ((I - K_t C_t)(x_t - \bar{\mu}_t))^T] \\ &\quad + E[(K_t \delta_t) (K_t \delta_t)^T]\end{aligned}$$

KALMAN FILTER

○ Measurement Update

- Recall multiplication by a constant yields

$$\begin{aligned}\text{cov}(Ax) &= E[(Ax - A\mu)(Ax - A\mu)^T] \\ &= A\text{cov}(X)A^T\end{aligned}$$

- In the above, numerous constants

$$\begin{aligned}\Sigma_t &= E\left[\left((I - K_t C_t)(x_t - \bar{\mu}_t)\right)\left((I - K_t C_t)(x_t - \bar{\mu}_t)\right)^T\right] \\ &\quad + E\left[\left(K_t \delta_t\right)\left(K_t \delta_t\right)^T\right] \\ &= \text{cov}\left(\left(I - K_t C_t\right)(x_t - \bar{\mu}_t)\right) + \text{cov}\left(K_t \delta_t\right)\end{aligned}$$

KALMAN FILTER

○ Measurement Update

- The resulting covariance is

$$\begin{aligned}\Sigma_t &= (I - K_t C_t) E \left[(x_t - \bar{\mu}_t)(x_t - \bar{\mu}_t)^T \right] (I - K_t C_t)^T \\ &\quad + K_t E \left[\delta_t \delta_t^T \right] K_t^T\end{aligned}$$

- The expectations that remain are known quantities

$$\begin{aligned}\Sigma_t &= (I - K_t C_t) \bar{\Sigma}_t (I - K_t C_t)^T + K_t Q_t K_t^T \\ &= \bar{\Sigma}_t - K_t C_t \bar{\Sigma}_t - \bar{\Sigma}_t C_t^T K_t^T + K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) K_t^T\end{aligned}$$

- Which leaves us with a quadratic equation in K_t

KALMAN FILTER

- Measurement update

- We now have the covariance in a form that can be optimized

$$\min \operatorname{tr}(\Sigma_t)$$

- We need two identities to find this minimum
 - Differentiation of linear matrix expression

$$\frac{\partial \operatorname{tr}(AXB)}{\partial X} = \frac{\partial \operatorname{tr}(B^T X^T A^T)}{\partial X} = A^T B^T$$

- Differentiation of quadratic matrix expression

$$\frac{\partial \operatorname{tr}(XAX^T)}{\partial X} = XA^T + XA$$

KALMAN FILTER

○ Measurement update

- The optimization is done by setting the derivative of the trace w.r.t the Kalman gain to 0

$$\min \operatorname{tr} \left(\bar{\Sigma}_t - K_t C_t \bar{\Sigma}_t - \bar{\Sigma}_t C_t^T K_t^T + K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) K_t^T \right)$$

- Taking the matrix derivative w.r.t K_t
 - Two linear terms and one quadratic

$$\begin{aligned} \frac{\partial}{\partial K_t} \operatorname{tr}(\Sigma_t) = & -\bar{\Sigma}_t^T C_t^T - \bar{\Sigma}_t C_t^T \\ & + K_t (C_t \bar{\Sigma}_t C_t^T + Q_t)^T + K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) \end{aligned}$$

KALMAN FILTER

○ Measurement update

- Set the derivative to 0, noting that covariance is symmetric, and AXA^T preserves symmetry

$$\frac{\partial}{\partial K_t} \text{tr}(\Sigma_t) = \left(-2\bar{\Sigma}_t C_t^T + 2K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) \right) = 0$$

- Simplifying

$$K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) = \bar{\Sigma}_t C_t^T$$

- And finally, we arrive at the Kalman gain equation

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

KALMAN FILTER

○ Measurement Update

- So far, we have found the optimal gain K_t which minimizes mean square error in the measurement update for the mean

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

$$\mu_t = \bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t)$$

- Next, we need to simplify the covariance update using this result for the Kalman gain

KALMAN FILTER

- Measurement Update

- Recall the Covariance update was

$$\Sigma_t = \bar{\Sigma}_t - K_t C_t \bar{\Sigma}_t - \bar{\Sigma}_t C_t^T K_t^T + K_t (C_t \bar{\Sigma}_t C_t^T + Q_t) K_t^T$$

- Substituting in the Kalman gain gives

$$\begin{aligned} \Sigma_t &= \bar{\Sigma}_t - \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1} C_t \bar{\Sigma}_t \\ &\quad - \bar{\Sigma}_t C_t^T \left(\bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1} \right)^T \\ &\quad + \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1} (C_t \bar{\Sigma}_t C_t^T + Q_t) \left(\bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1} \right)^T \end{aligned}$$

KALMAN FILTER

○ Measurement Update

- Fortunately, almost everything cancels and we are left with

$$\begin{aligned}\Sigma_t &= \bar{\Sigma}_t - \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1} C_t \bar{\Sigma}_t \\ &= (I - K_t C_t) \bar{\Sigma}_t\end{aligned}$$

KALMAN FILTER

○ Kalman Filter Algorithm

- At each time step, t , update both sets of beliefs
 1. Prediction update

$$\bar{\mu}_t = A_t \mu_{t-1} + B_t u_t$$

$$\bar{\Sigma}_t = A_t \Sigma_{t-1} A_t^T + R_t$$

2. Measurement update

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

$$\mu_t = \bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t)$$

$$\Sigma_t = (I - K_t C_t) \bar{\Sigma}_t$$

- Kalman Gain, K_t
 - Blending factor between prediction and measurement

KALMAN FILTER

○ Summary

- Follows same framework as Bayes filter
- Requires linear motion and Gaussian disturbance
- Requires linear measurement and Gaussian noise
- It is sufficient to update mean and covariance of beliefs, because they remain Gaussian
- Prediction step involves addition of Gaussians
- Measurement step seeks to minimize mean square error of the estimate
 - Expand out covariance from definition and measurement model
 - Assume form of estimator, linear combination of prediction and measurement
 - Solve MMSE problem to find optimal linear combination
 - Simplify covariance update once gain is found

KALMAN FILTERING

○ Relation to Bayes Filter Problem Formulation

- State prior

$$bel(x_0) = p(x_0) = N(\mu_0, \Sigma_0)$$

- Motion model

$$p(x_t | x_{t-1}, u_t) = N(A_t x_{t-1} + B_t u_t, A_t \Sigma_{t-1} A_t^T + R_t)$$

- Measurement model

$$p(y_t | x_t) = N(C_t x_t, Q_t)$$

- Beliefs

$$bel(x_t) = N(\mu_t, \Sigma_t), \quad \overline{bel}(x_t) = N(\bar{\mu}_t, \bar{\Sigma}_t)$$

KALMAN FILTER

○ Relation to Bayes Filter Algorithm

1. Prediction update (Total probability)
 - Insert normal distributions

$$\begin{aligned}\overline{bel}(x_t) &= \int p(x_t | u_t, x_{t-1}) bel(x_{t-1}) dx_{t-1} \\ &= \eta \int e^{-1/2(x_t - A_t x_{t-1} - B_t u_t)^T R_t^{-1} (x_t - A_t x_{t-1} - B_t u_t)} e^{-1/2(x_{t-1} - \mu_{t-1})^T \Sigma_{t-1}^{-1} (x_{t-1} - \mu_{t-1})} dx_{t-1}\end{aligned}$$

- Separate out terms that depend on current state
- Manipulate remaining integral into a Gaussian pdf form of previous state
- Integrate over full range to get 1
- Manipulate remaining terms and solve for Kalman prediction equations.

$$\sim N(A_t \mu_{t-1} + B_t u_t, A_t \Sigma_{t-1} A_t^T + R_t)$$

- Refer to Thrun, Burgard & Fox Chap. 3 for details

KALMAN FILTER

- Relation to Bayes Filter Algorithm

- 2. Measurement update (Bayes Theorem)

$$\begin{aligned} \text{bel}(x_t) &= \eta p(y_t | x_t) \overline{\text{bel}}(x_t) \\ &= \eta e^{-1/2(y_t - C_t x_t)^T Q_t^{-1} (y_t - C_t x_t)} e^{-1/2(x_t - \bar{\mu}_t)^T \bar{\Sigma}_t^{-1} (x_t - \bar{\mu}_t)} \end{aligned}$$

- Reorganize exponents and note it remains a Gaussian
- For any Gaussian:
 - Second derivative of exponent is inverse of covariance
 - Mean minimizes exponent,
 - Set first derivative of exponent to 0 and solve
- Use this to solve for mean and covariance of belief

$$= N(\bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t), (I - K_t C_t) \bar{\Sigma}_t)$$

- where K_t is the Kalman gain as before

KALMAN FILTER

○ Example

- 3D Linear motion model for three thruster AUV (heading constant)

- State $x = \begin{bmatrix} p_n \\ v_n \\ p_e \\ v_e \\ p_d \\ v_d \end{bmatrix}$ Input $u = \begin{bmatrix} T_n \\ T_e \\ T_d \end{bmatrix}$

- Continuous dynamics for

$$\dot{x}_n(t) = v_n(t)$$

$$m\dot{v}_n(t) = -bv_n(t) + T_n(t)$$



KALMAN FILTER



○ Example – Omni-directional AUV

- Discrete Dynamics from zero order hold, $dt = 0.1s$

$$x_t = \begin{bmatrix} 1 & .0975 & 0 & 0 \\ 0 & .9512 & 0 & 0 \\ 0 & 0 & 1 & .0975 \\ 0 & 0 & 0 & .9512 \end{bmatrix} x_{t-1} + \begin{bmatrix} .0025 & 0 \\ .0488 & 0 \\ 0 & .0025 \\ 0 & .0048 \end{bmatrix} u_t + \varepsilon_t$$

- Disturbances

$$\varepsilon_t \sim N \left(0, R = \begin{bmatrix} 0.01 & 0 & 0 & 0 \\ 0 & 0.01 & 0 & 0 \\ 0 & 0 & 0.01 & 0 \\ 0 & 0 & 0 & 0.01 \end{bmatrix} \right)$$

KALMAN FILTER



○ Example – Omni-directional AUV

- Measurement Model
 - Can only measure position (relative to known objects)

$$y_t = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{bmatrix} x_t + \delta_t$$

- With correlated measurement noise

$$\delta_t \sim N\left(0, Q = \begin{bmatrix} 0.4 & -0.1 \\ -0.1 & 0.1 \end{bmatrix}\right)$$

KALMAN FILTER



○ Example

- Control inputs

- This time different frequencies of sinusoidal input

$$u_t = 10 \begin{bmatrix} \sin(2t) \\ \cos(t) \end{bmatrix}$$

- Simulation run for 10 seconds, or 101 time steps

- Prior distribution

- Fairly course initialization

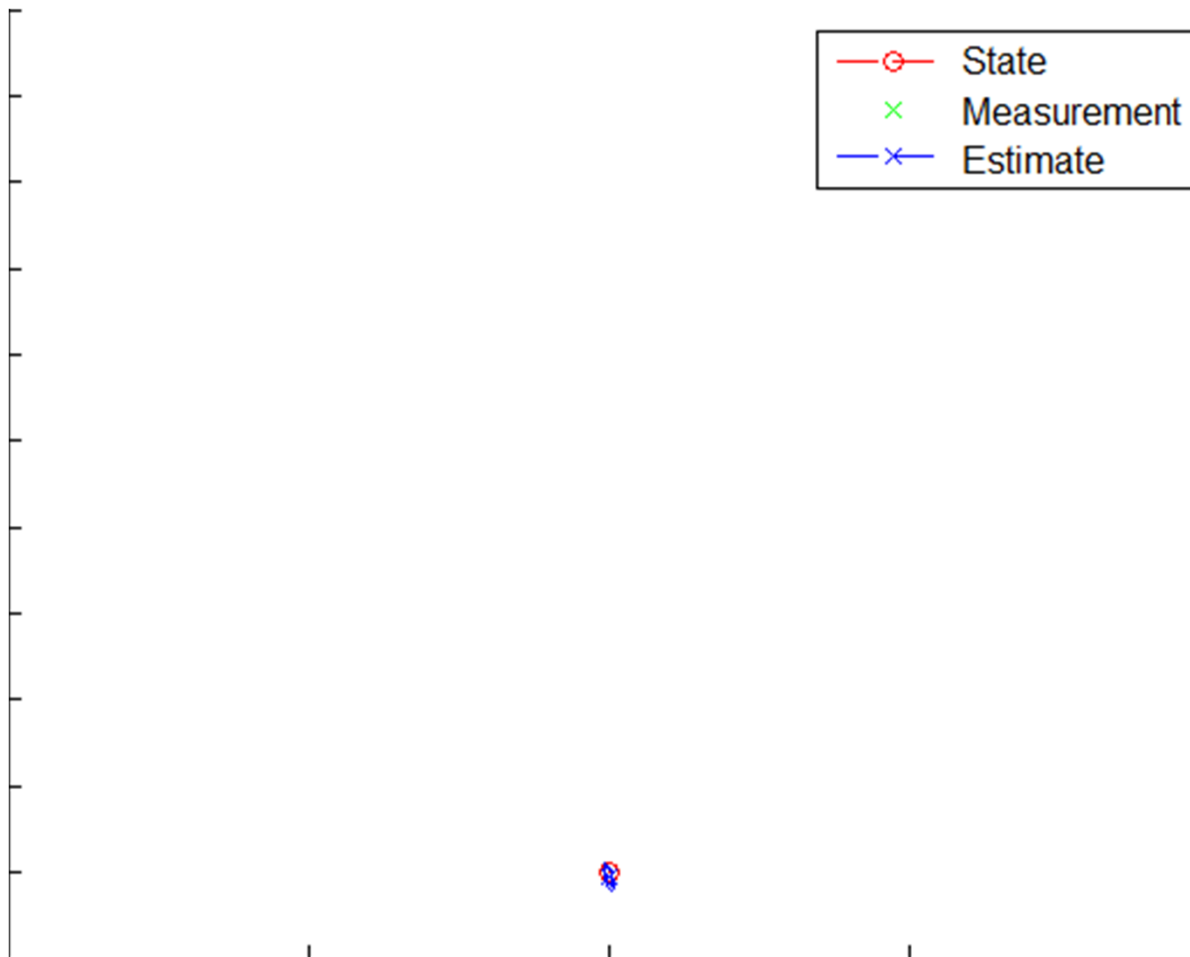
$$bel(x_0) \sim N(0, I)$$

KALMAN FILTER



○ Example

- Ideal results: very low noise and disturbance levels

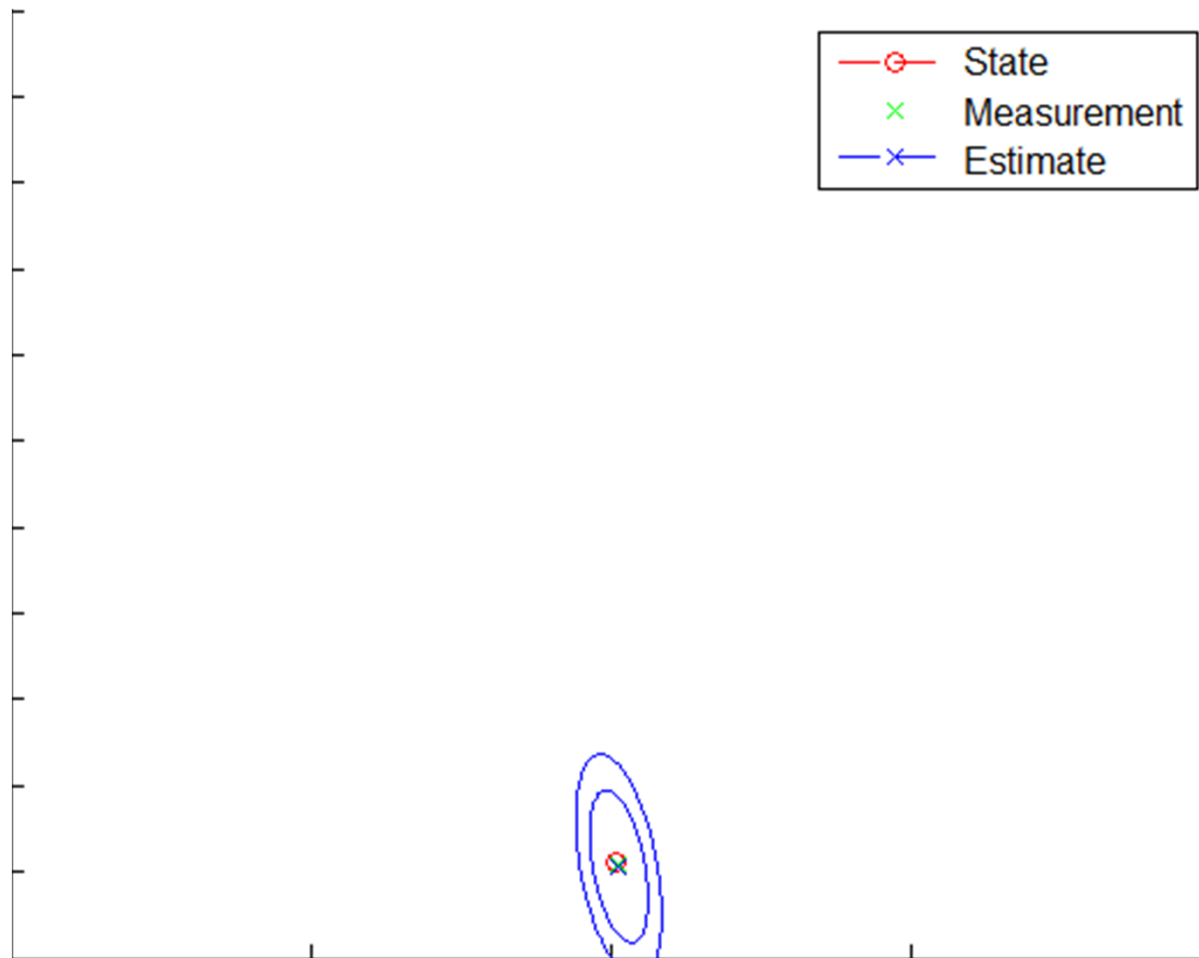


KALMAN FILTER



○ Example

- Results with larger noise and disturbances

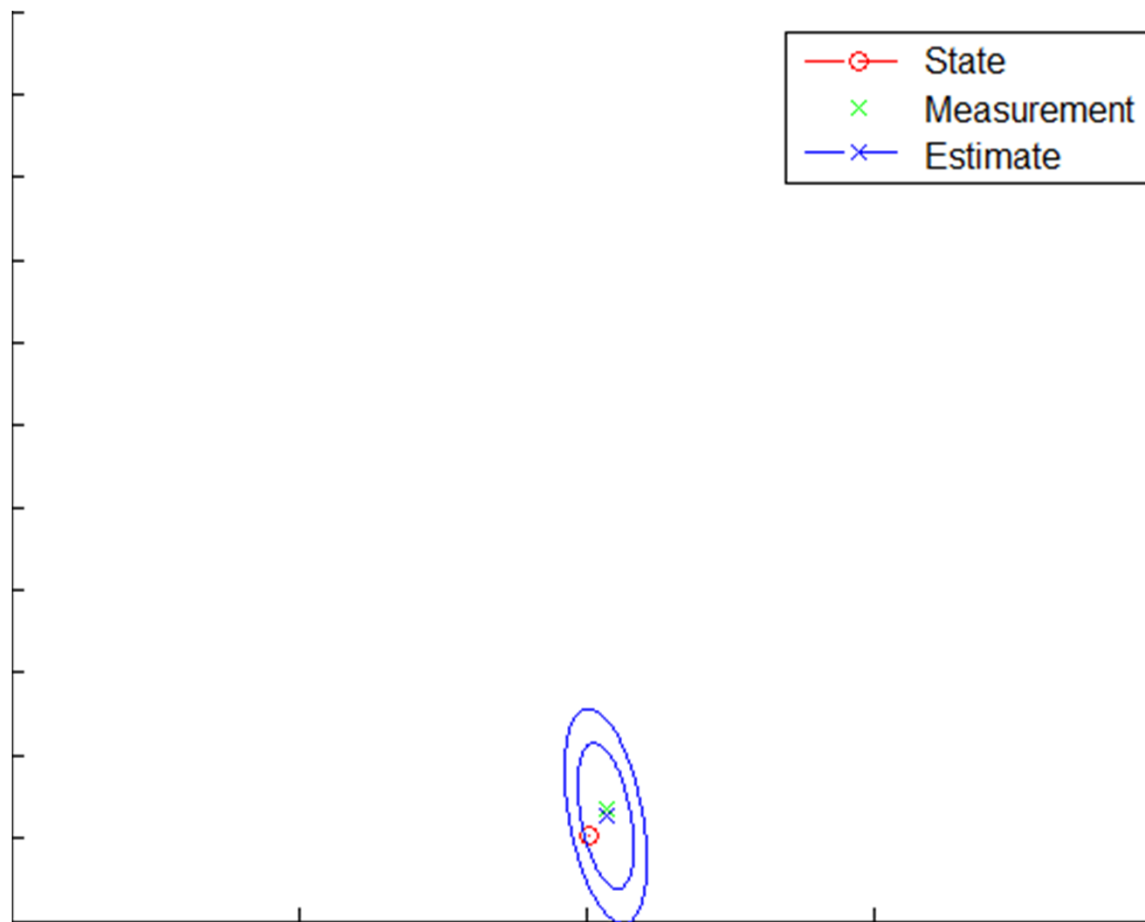


KALMAN FILTER



○ Example

- Effect of decreased motion disturbances

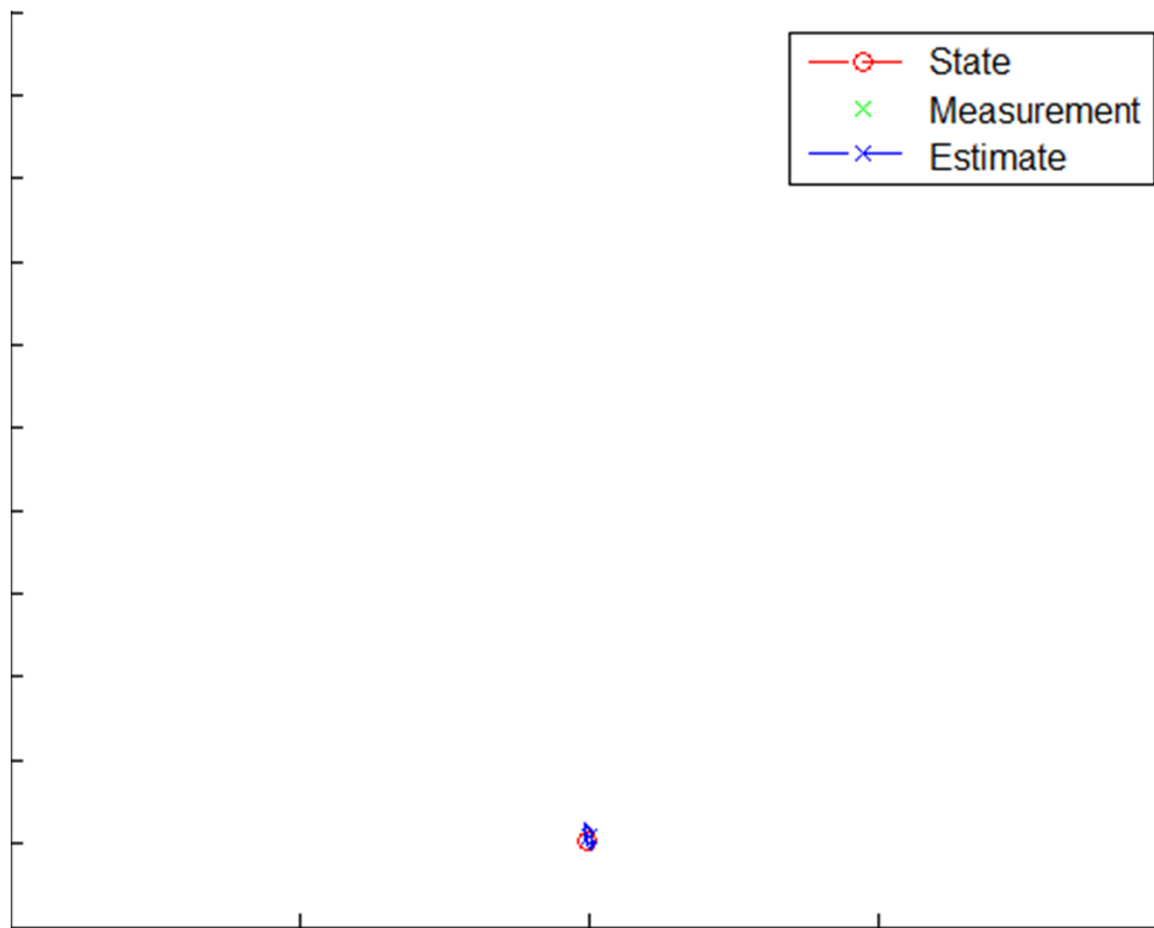


KALMAN FILTER



○ Example

- Effect of decreased measurement noise



KALMAN FILTER

- Belief mean is tradeoff between prediction and measurement

$$\mu_t = \bar{\mu}_t + K_t(y_t - C_t\bar{\mu}_t)$$

- Kalman gain determines how to blend estimates

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

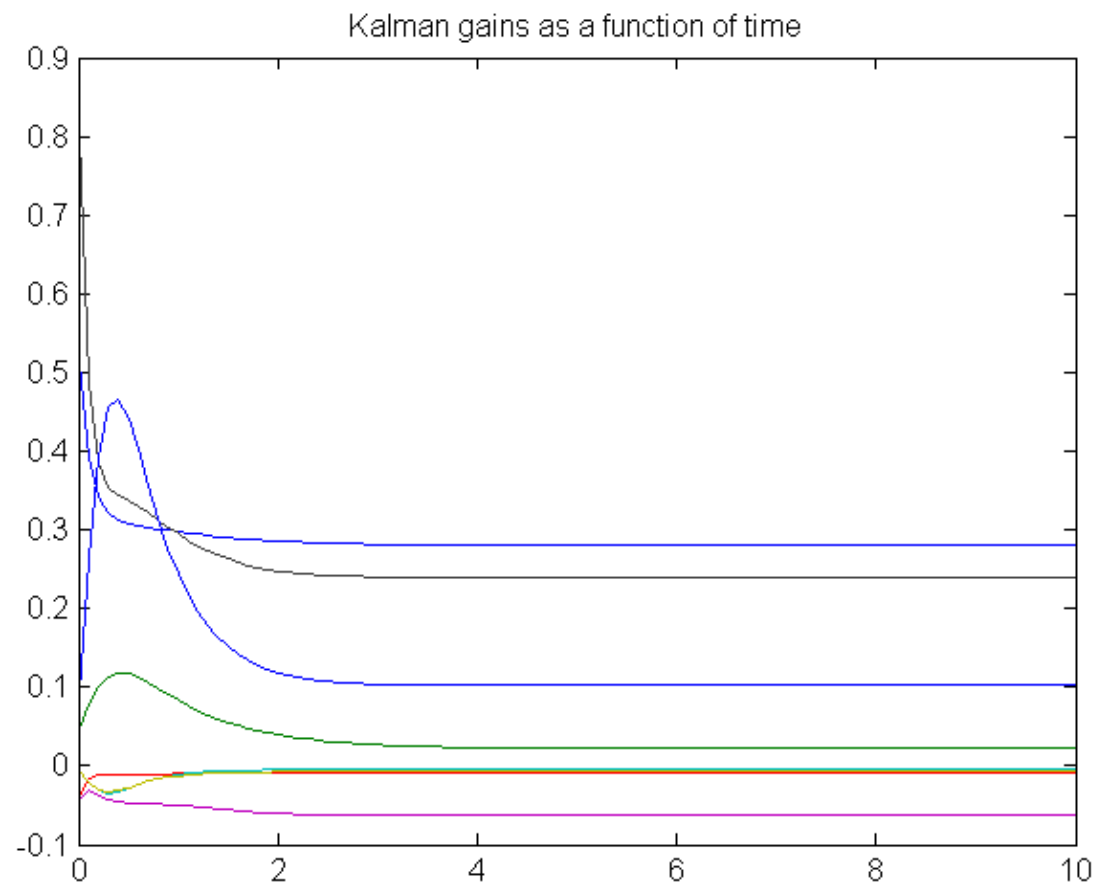
- If Q_t is large, inverse is small, so Kalman gain remains small
 - When measurements are high in covariance, don't trust them!
- If R_t is large, then so is predicted belief covariance, so Kalman gain becomes large
 - When model is affected by large unknown disturbances, don't trust the predicted motion!

KALMAN FILTER



○ Example

- Evolution of Kalman Gain



KALMAN FILTER

○ Steady state Kalman Filter

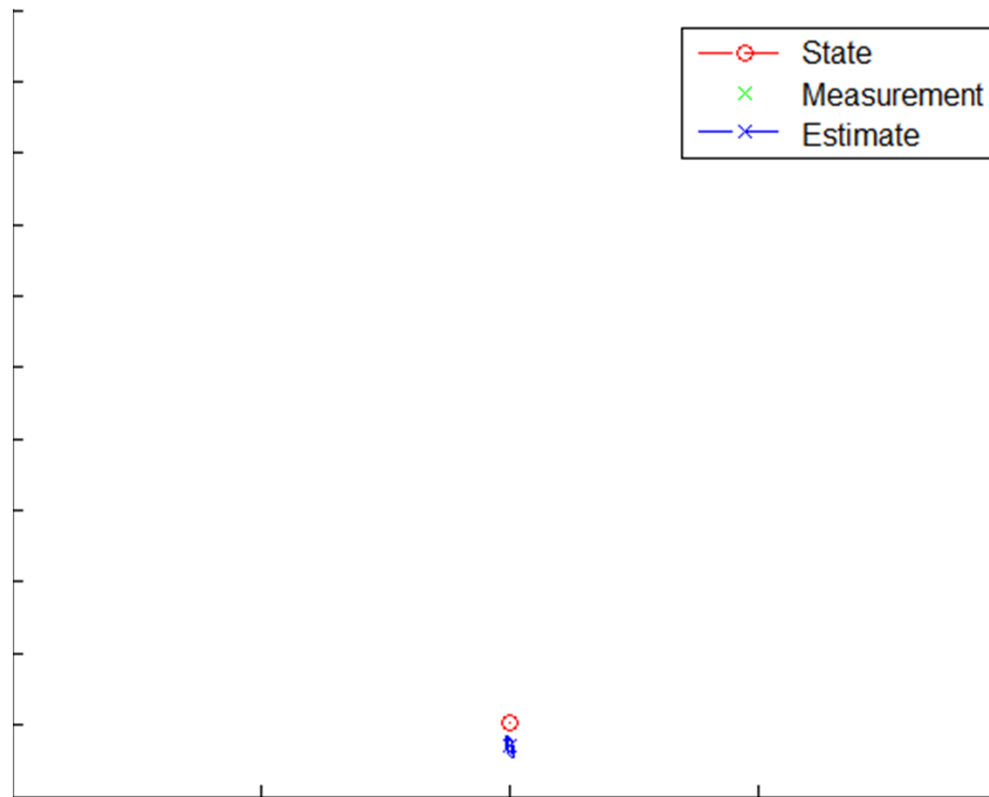
- For constant noise/disturbance models, it is possible to use steady state values for the Kalman gain
 - Set $\Sigma = \Sigma_t = \Sigma_{t-1}$ in the Kalman filter update equations and solve for Σ
 - Referred to as the Discrete Algebraic Ricatti Equation (DARE), Matlab will solve it for you
- Can also run Kalman filter until convergence and then eliminate gain update step (matrix inversion)

KALMAN FILTER



○ Example

- Incorrect measurement distribution (covariance actually much larger)



- Estimate tracks measurements too closely

KALMAN FILTER

○ Multi-Rate Kalman Filter

- At each time step, it is possible to use different measurement models
 - Time varying C_t and Q_t
- Identify a base update rate
 - Find greatest common divisor of sample rates
 - e.g. GPS 5Hz, Sodar 12 Hz, Base rate 60 Hz
- Create discretized motion model at base rate
- At each timestep
 1. Perform prediction update
 2. If new measurements exist, perform measurement update for those measurements only
 - Select appropriate C_t and Q_t

KALMAN FILTER



○ Example

• Multi-rate Kalman Filter

- Add in velocity measurements at 100 Hz
- Base update rate 0.01 s
- Create two separate types of measurement updates
 - Velocity only measurement for 9 time steps

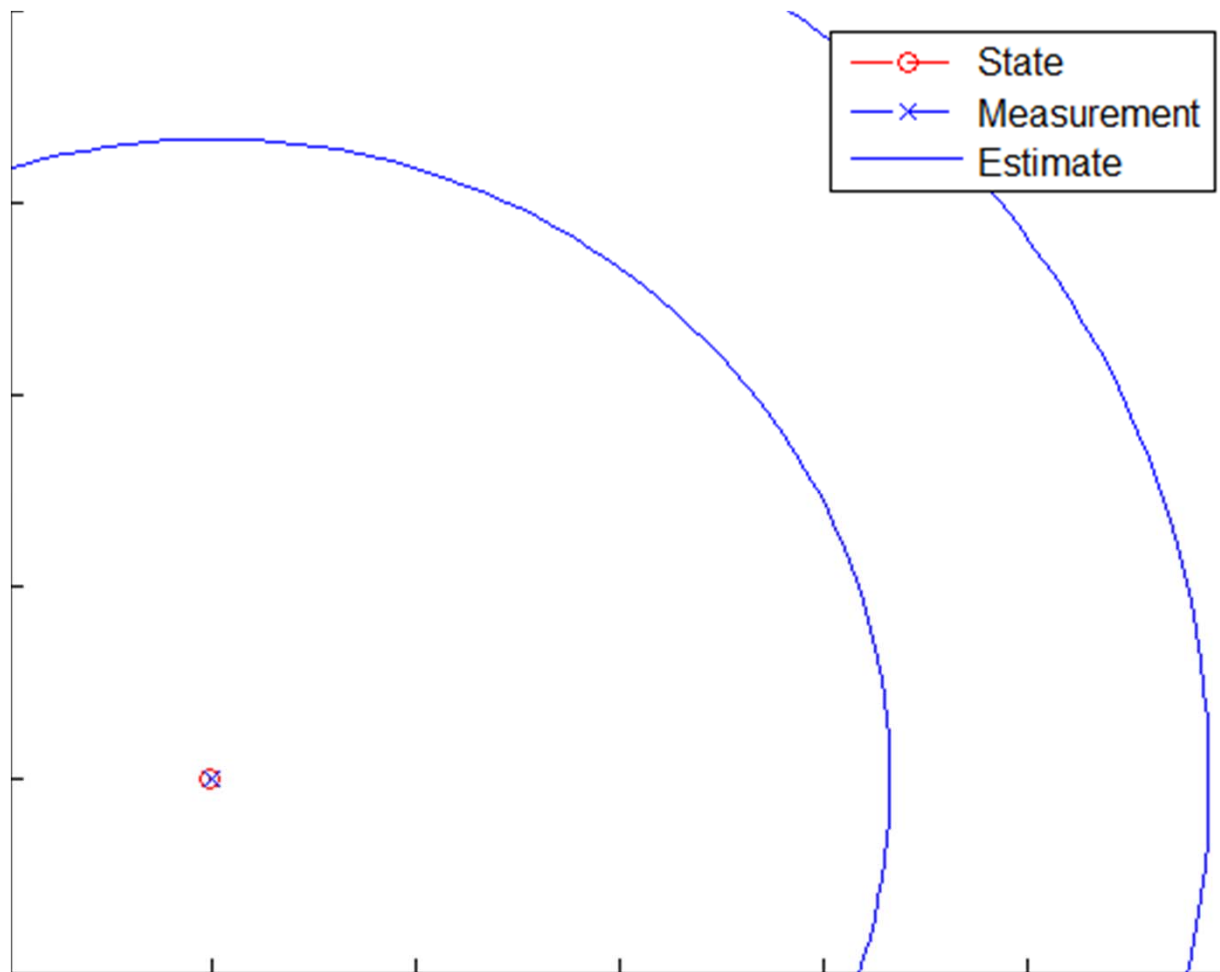
$$C_{1:9} = \begin{bmatrix} 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \quad Q_{1:9} = \begin{bmatrix} 0.1 & -0.01 \\ -0.01 & 0.05 \end{bmatrix}$$

- Full state measurement on the 10th time step

$$C_{10} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \quad Q_{10} = \begin{bmatrix} 0.004 & 0 & -0.001 & 0 \\ 0 & 0.1 & 0 & -0.01 \\ -0.001 & 0 & 0.001 & 0 \\ 0 & -0.01 & 0 & 0.05 \end{bmatrix}$$

KALMAN FILTER

- Example
 - Multi-rate estimation



KALMAN FILTER

- Alternate formulation: Information Filter
 - Provides possibility for computational savings when taking many redundant measurements
 - Based on information theory concepts (Fisher Information)
 - Define the Information Matrix as the inverse of the covariance

$$\Omega_t = \Sigma_t^{-1}$$

- Define the Information vector as

$$\xi_t = \Sigma_t^{-1} \mu_t$$

KALMAN FILTER

○ Information Filter

- Substitution into the Kalman filter equation yields
 1. Prediction update

$$\bar{\xi}_t = \bar{\Omega}_t (A_t \Omega_{t-1}^{-1} \xi_{t-1} + B_t u_t)$$

$$\bar{\Omega}_t = (A_t \Omega_{t-1}^{-1} A_t^T + R_t)^{-1}$$

2. Measurement update

$$\xi_t = C_t Q_t^{-1} y_t + \bar{\xi}_t$$

$$\Omega_t = C_t^T Q_t^{-1} C_t + \bar{\Omega}_t$$

KALMAN FILTER

○ Information Filter

- The matrix inversion is now embedded in the prediction update
 - Belief and predicted belief inverse depend on the number of states
 - For Kalman filter, gain inverse depends on the number of measurements
 - This can be a significant savings in some cases
- To compute state estimate

$$\mu_t = \Omega_t^{-1} \xi_t$$

- Covariance already calculated

OUTLINE

- Bayes Filter Framework
- Kalman Filter
- Extended Kalman Filter
- Particle Filter

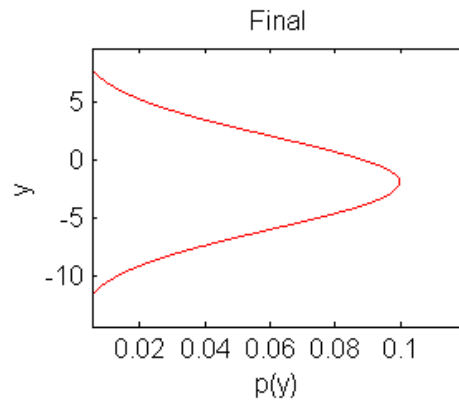
EXTENDED KALMAN FILTER

- Kalman Filter requires linear motion and measurement models
 - Results in compact, recursive estimation technique
 - Not very realistic for most applications
- Nonlinear models eliminate the guarantee that the belief distributions remain Gaussian
 - No longer able to simply track mean and covariance
 - No closed form solution to Bayes filter algorithm can be found for general nonlinear model

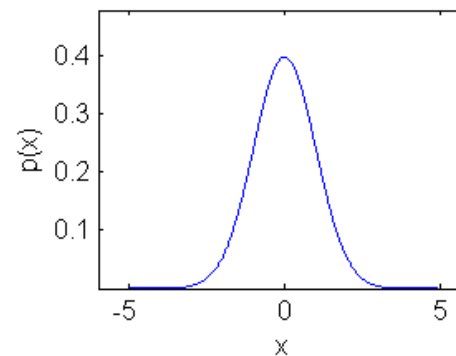
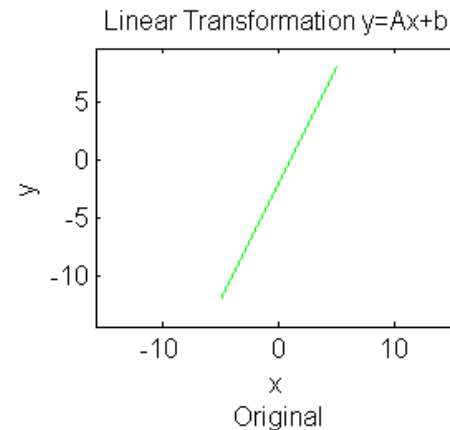
EXTENDED KALMAN FILTER

- Effect of nonlinearity on Gaussian distribution
 - Linear transformation

$$y = 2x - 2$$



$$y \sim N(-2, 4)$$



$$x \sim N(0, 1)$$

EXTENDED KALMAN FILTER

- Arbitrary distribution generation

- Take 5,000,000 samples of original Gaussian

$$x^i \sim N(0,1), \quad i = 1, \dots, n$$

- Apply nonlinear transformation to each sample

$$y^i = \tan^{-1}(x^i)$$

- Create histogram with 100 bins and normalize counts

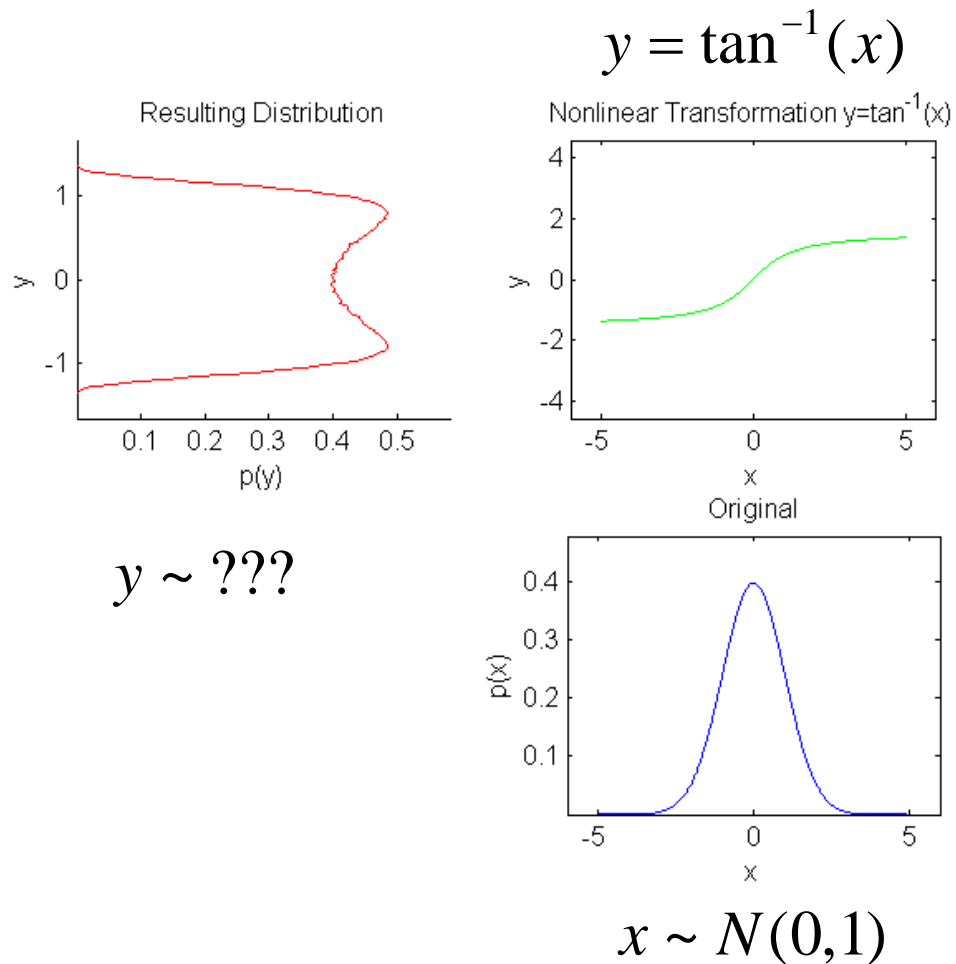
- Best Gaussian fit generation

- Calculate mean and covariance of 5,000,000 transformed samples

$$\mu_{BG} = \frac{1}{n} \sum_{i=1}^n y^i \quad \Sigma_{BG} = \frac{1}{n-1} \sum_{i=1}^n (y^i - \mu_{BG})(y^i - \mu_{BG})^T$$

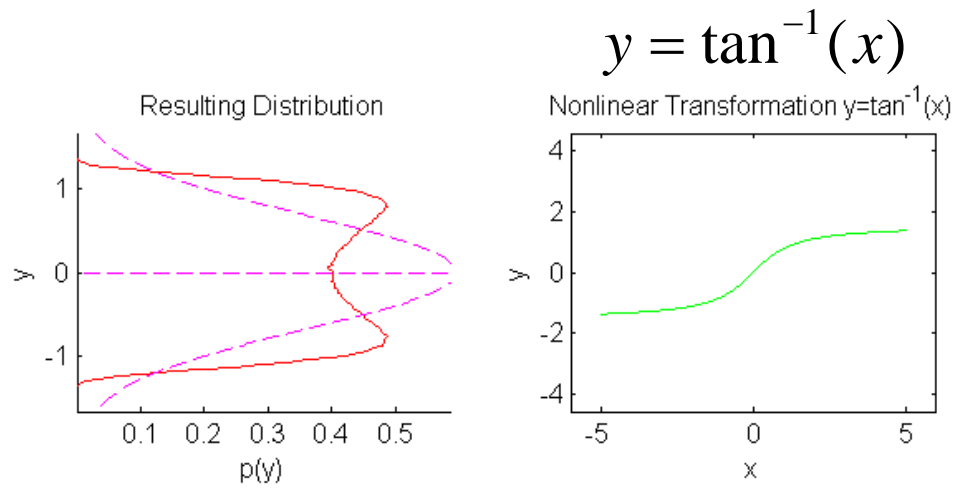
EXTENDED KALMAN FILTER

- Effect of nonlinearity on Gaussian distribution
 - Nonlinear transformation

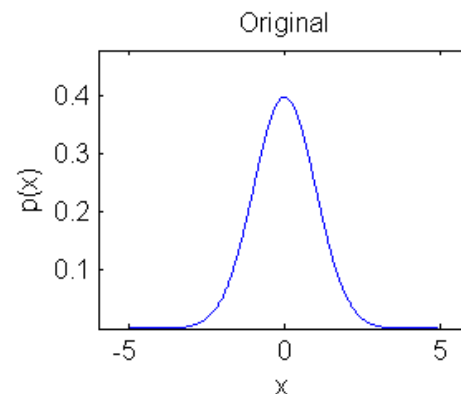


EXTENDED KALMAN FILTER

- Effect of nonlinearity on Gaussian distribution
 - Nonlinear transformation

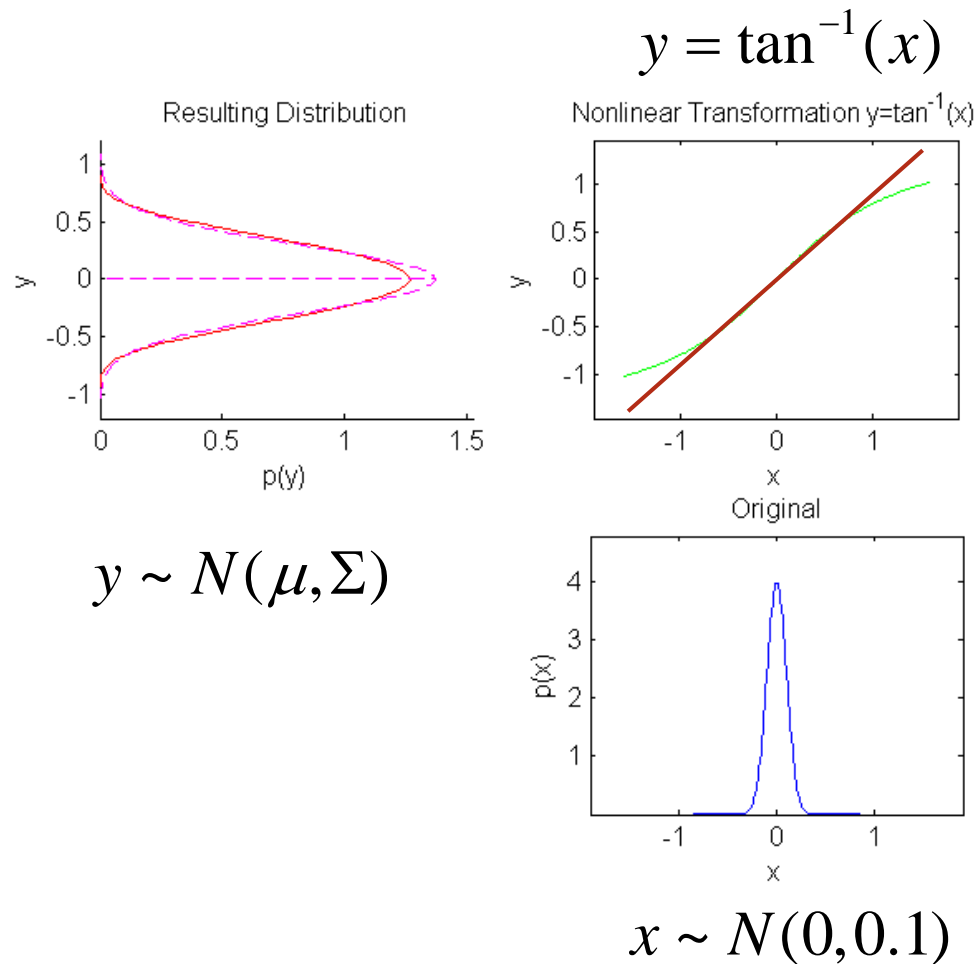


$y \sim N(\mu, \Sigma)??$



EXTENDED KALMAN FILTER

- Effect of nonlinearity on Gaussian distribution
 - Nonlinear transformation



EXTENDED KALMAN FILTER

○ Extended Kalman Filter

- A direct generalization of the Kalman filter to nonlinear motion and measurement models
 - Relies on linearization about current estimate
 - Works well when the problem maintains locally linear and Gaussian characteristics
 - Computationally similar to Kalman Filter
 - Covariance can diverge when approximation is poor!

EXTENDED KALMAN FILTER

- Extended Kalman Filter Modeling Assumption

- Prior over the state is Gaussian

$$p(x_0) \sim N(\mu_0, \Sigma_0)$$

- Motion model, nonlinear but still with additive Gaussian disturbances

$$x_t = g(x_{t-1}, u_t) + \varepsilon_t \quad \varepsilon_t \sim N(0, R_t)$$

- Measurement model also nonlinear with additive Gaussian noise

$$y_t = h(x_t) + \delta_t \quad \delta_t \sim N(0, Q_t)$$

- Nonlinearity destroys certainty that beliefs remain Gaussian

EXTENDED KALMAN FILTER

○ Recall Kalman Filter Algorithm

1. Prediction update

$$\bar{\mu}_t = A_t \mu_{t-1} + B_t u_t$$

$$\bar{\Sigma}_t = A_t \Sigma_{t-1} A_t^T + R_t$$

2. Measurement update

$$K_t = \bar{\Sigma}_t C_t^T (C_t \bar{\Sigma}_t C_t^T + Q_t)^{-1}$$

$$\mu_t = \bar{\mu}_t + K_t (y_t - C_t \bar{\mu}_t)$$

$$\Sigma_t = (I - K_t C_t) \bar{\Sigma}_t$$

- B_t only enters predicted mean calculation
- A_t, C_t also affect covariance

EXTENDED KALMAN FILTER

- How to update beliefs while maintaining Gaussian form of distribution?
 - Key idea of EKF
 - The mean can be propagated through the nonlinear model
 - The covariance can be updated with a locally linear approximation to the model

EXTENDED KALMAN FILTER

- First Order Taylor Series Expansion

- Motion model
 - Linearize about most likely state (the previous mean)

$$\begin{aligned} g(x_{t-1}, u_t) &\approx g(\mu_{t-1}, u_t) + \frac{\partial}{\partial x_{t-1}} g(x_{t-1}, u_t) \Big|_{x_{t-1}=\mu_{t-1}} (x_{t-1} - \mu_{t-1}) \\ &= g(\mu_{t-1}, u_t) + G_t \cdot (x_{t-1} - \mu_{t-1}) \end{aligned}$$

EXTENDED KALMAN FILTER

- First Order Taylor Series Expansion

- Measurement Model

- Linearize about most likely state (the predicted mean)

$$\begin{aligned}h(x_t) &\approx h(\bar{\mu}_t) + \left. \frac{\partial}{\partial x_t} h(x_t) \right|_{x_t = \bar{\mu}_t} (x_t - \bar{\mu}_t) \\ &= h(\bar{\mu}_t) + H_t \cdot (x_t - \bar{\mu}_t)\end{aligned}$$

- Both models are now linear

- Only valid near point of linearization

EXTENDED KALMAN FILTER

○ Prediction Update

- Only new information is input u_t
- Prediction update is a linear transformation of belief at previous time step
 - Motion model is

$$x_t = g(\mu_{t-1}, u_t) + G_t \cdot (x_{t-1} - \mu_{t-1}) + \varepsilon_t$$

- Motion disturbance, previous belief are Gaussian so this is remains addition of Gaussian distributions

$$bel(x_{t-1}) \sim N(\mu_{t-1}, \Sigma_{t-1}) \quad \varepsilon_t \sim N(0, R_t)$$

- Therefore the predicted mean and covariance are

$$\bar{\mu}_t = g(\mu_{t-1}, u_t) + 0 + 0$$

$$\bar{\Sigma}_t = 0 + G_t \Sigma_{t-1} G_t^T + R_t$$

EXTENDED KALMAN FILTER

○ Measurement Update

- Follows same arguments as Kalman Filter derivation
 - MMSE estimator
 - Assume form of measurement update (linear, Kalman Gain)
 - Substitute in approximate measurement and motion models
 - Mean update relies on nonlinear model
 - Gain, covariance update rely on linearization

$$K_t = \bar{\Sigma}_t H_t^T (H_t \bar{\Sigma}_t H_t^T + Q_t)^{-1}$$

$$\mu_t = \bar{\mu}_t + K_t (y_t - h(\bar{\mu}_t))$$

$$\Sigma_t = (I - K_t H_t) \bar{\Sigma}_t$$

EXTENDED KALMAN FILTER

Extended Kalman Filter Algorithm

1. Prediction Update

$$G_t = \left. \frac{\partial}{\partial x_{t-1}} g(x_{t-1}, u_t) \right|_{x_{t-1} = \mu_{t-1}}$$

$$\bar{\mu}_t = g(\mu_{t-1}, u_t)$$

$$\bar{\Sigma}_t = G_t \Sigma_{t-1} G_t^T + R_t$$

2. Measurement Update

$$H_t = \left. \frac{\partial}{\partial x_t} h(x_t) \right|_{x_t = \bar{\mu}_t}$$

$$K_t = \bar{\Sigma}_t H_t^T (H_t \bar{\Sigma}_t H_t^T + Q_t)^{-1}$$

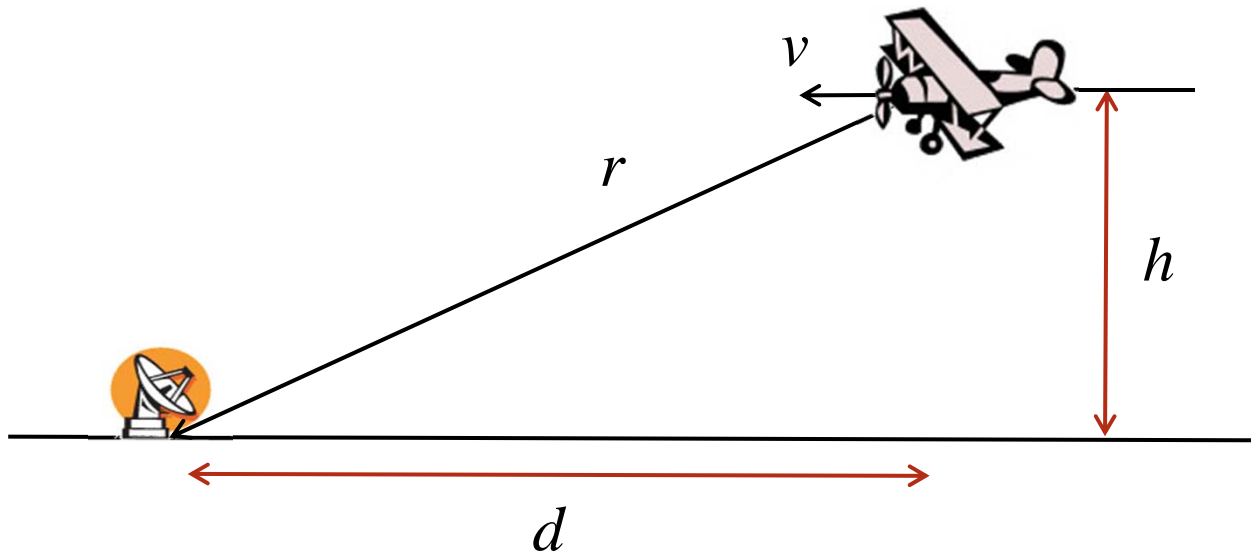
$$\mu_t = \bar{\mu}_t + K_t (y_t - h(\bar{\mu}_t))$$

$$\Sigma_t = (I - K_t H_t) \bar{\Sigma}_t$$

EXTENDED KALMAN FILTER

○ Example

- Radar measurement of an airplane position while flying at constant altitude and velocity



EXTENDED KALMAN FILTER

○ Example

- State

$$x = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} d \\ v \\ h \end{bmatrix}$$

Initial

$$x_0 = \begin{bmatrix} 20 \\ -2 \\ 3 \end{bmatrix}$$

- Motion Model

- Linear, no input (very simple)

$$x_t = g(x_{t-1}) + \varepsilon_t$$



$$x_{1,t} = x_{1,t-1} + x_{2,t-1} dt$$

$$x_{2,t} = x_{2,t-1}$$

$$x_{3,t} = x_{3,t-1}$$

EXTENDED KALMAN FILTER

○ Example

- Measurement Model

$$r = \sqrt{d^2 + h^2} + \delta_t$$

- Using state variables

$$y_t = \sqrt{x_{1,t}^2 + x_{3,t}^2} + \delta_t \quad h(x_t) = \sqrt{x_{1,t}^2 + x_{3,t}^2}$$

- Linearization of measurement model

$$\frac{\partial h}{\partial x_1} = \frac{1}{2} (x_1^2 + x_3^2)^{-1/2} 2x_1 = \frac{x_1}{\sqrt{x_1^2 + x_3^2}}$$

$$\frac{\partial h}{\partial x_3} = \frac{1}{2} (x_1^2 + x_3^2)^{-1/2} 2x_3 = \frac{x_3}{\sqrt{x_1^2 + x_3^2}}$$

EXTENDED KALMAN FILTER

○ Sample Code

```
%% Extended Kalman Filter Estimation
% Prediction update
mup = Ad*mu;
Sp = Ad*S*Ad' + R;

% Measurement update
Ht = [(mup(1))/(sqrt(mup(1)^2 + mup(3)^2));
      0;
      (mup(3))/(sqrt(mup(1)^2 + mup(3)^2))];

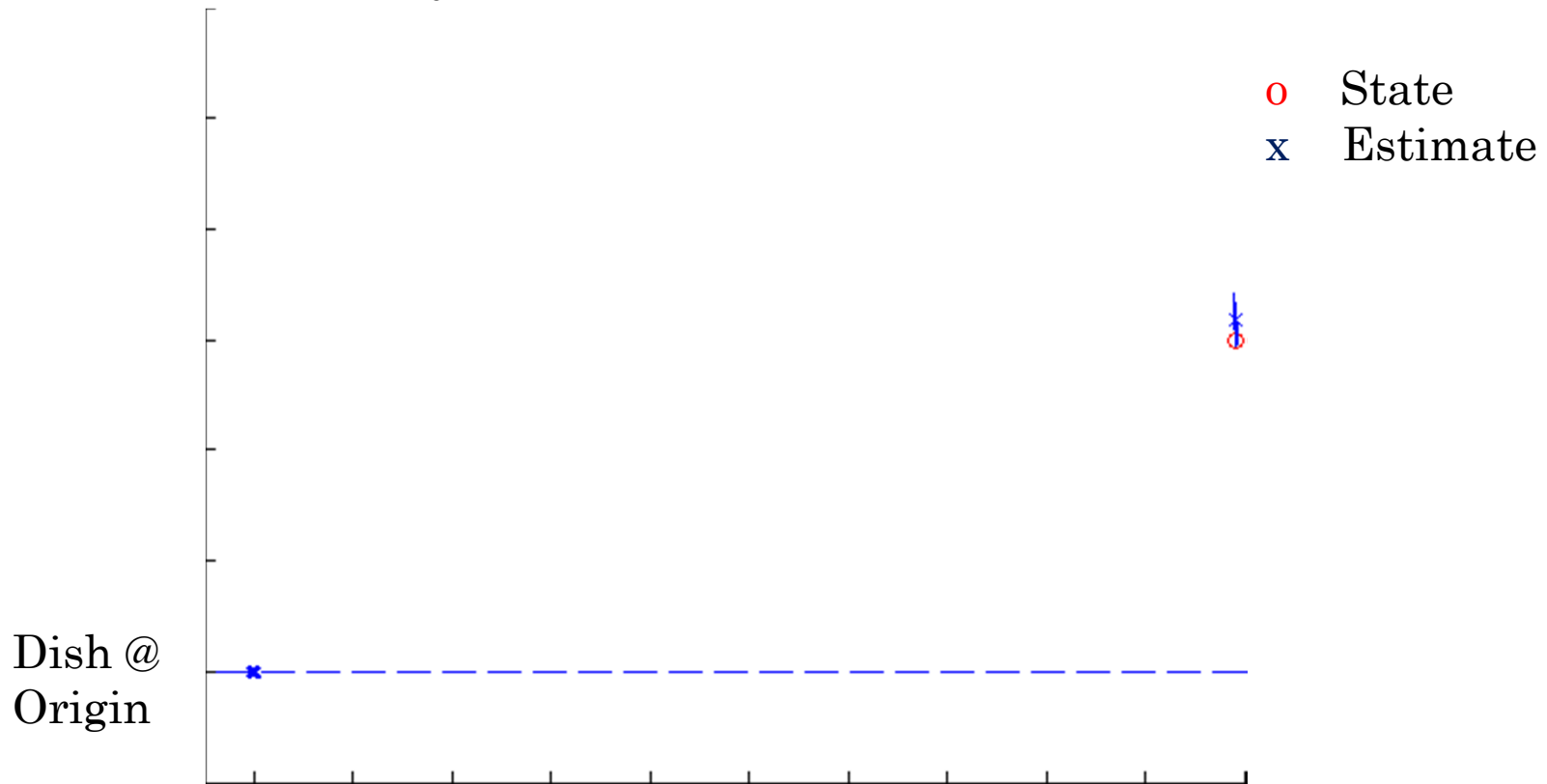
K = Sp*Ht'*inv(Ht*Sp*Ht'+Q);
mu = mup + K*(y(:,t)-sqrt(mup(1)^2 + mup(3)^2));
S = (eye(n)-K*Ht)*Sp;
```

EXTENDED KALMAN FILTER

○ Results

- Low noise, fairly accurate prior

$$\mu_0 = [22 \quad -1.8 \quad 3.5]^T$$

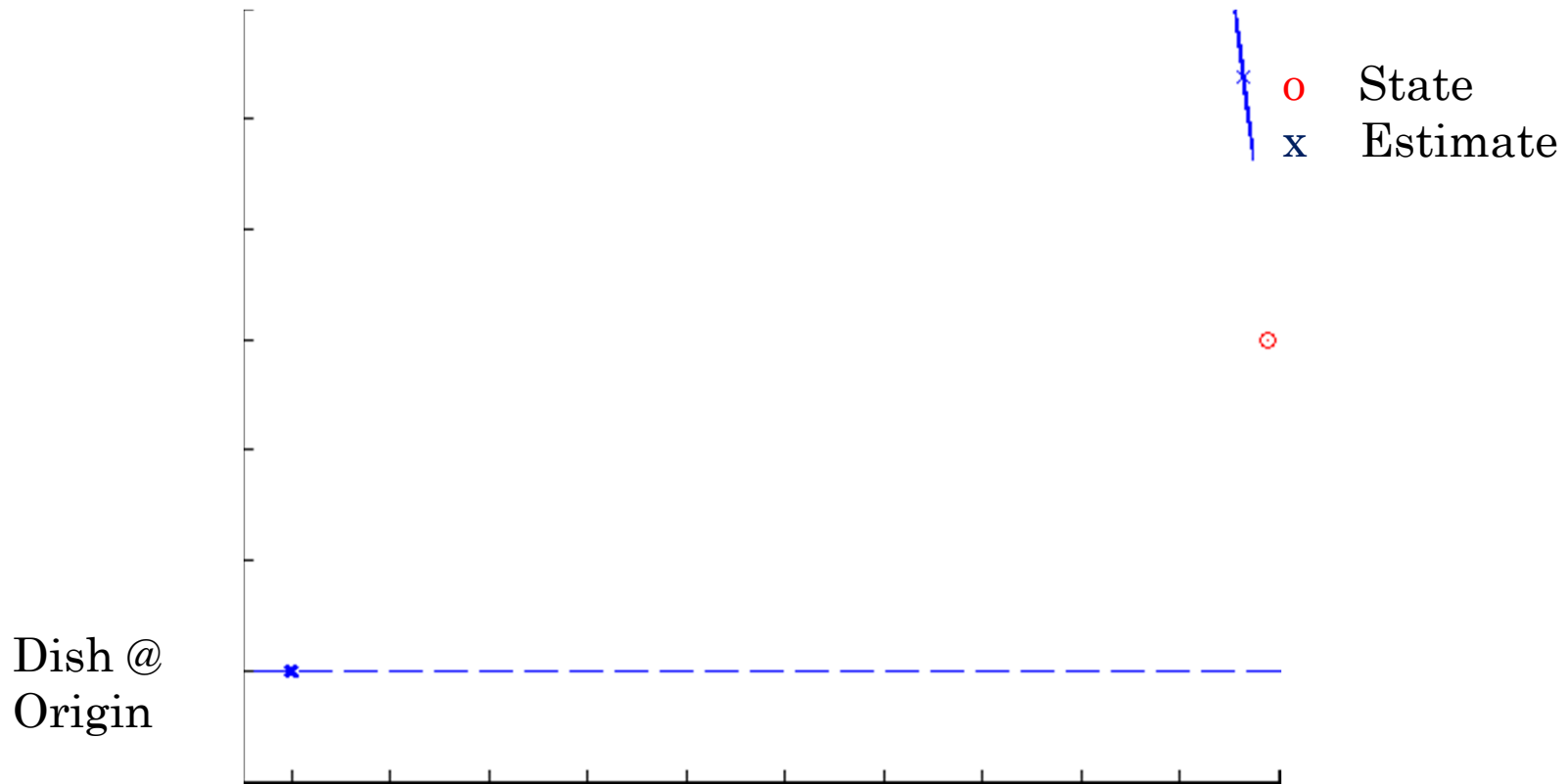


EXTENDED KALMAN FILTER

○ Results

- Low noise, incorrect prior

$$\mu_0 = [22 \quad -1.8 \quad 6]^T$$

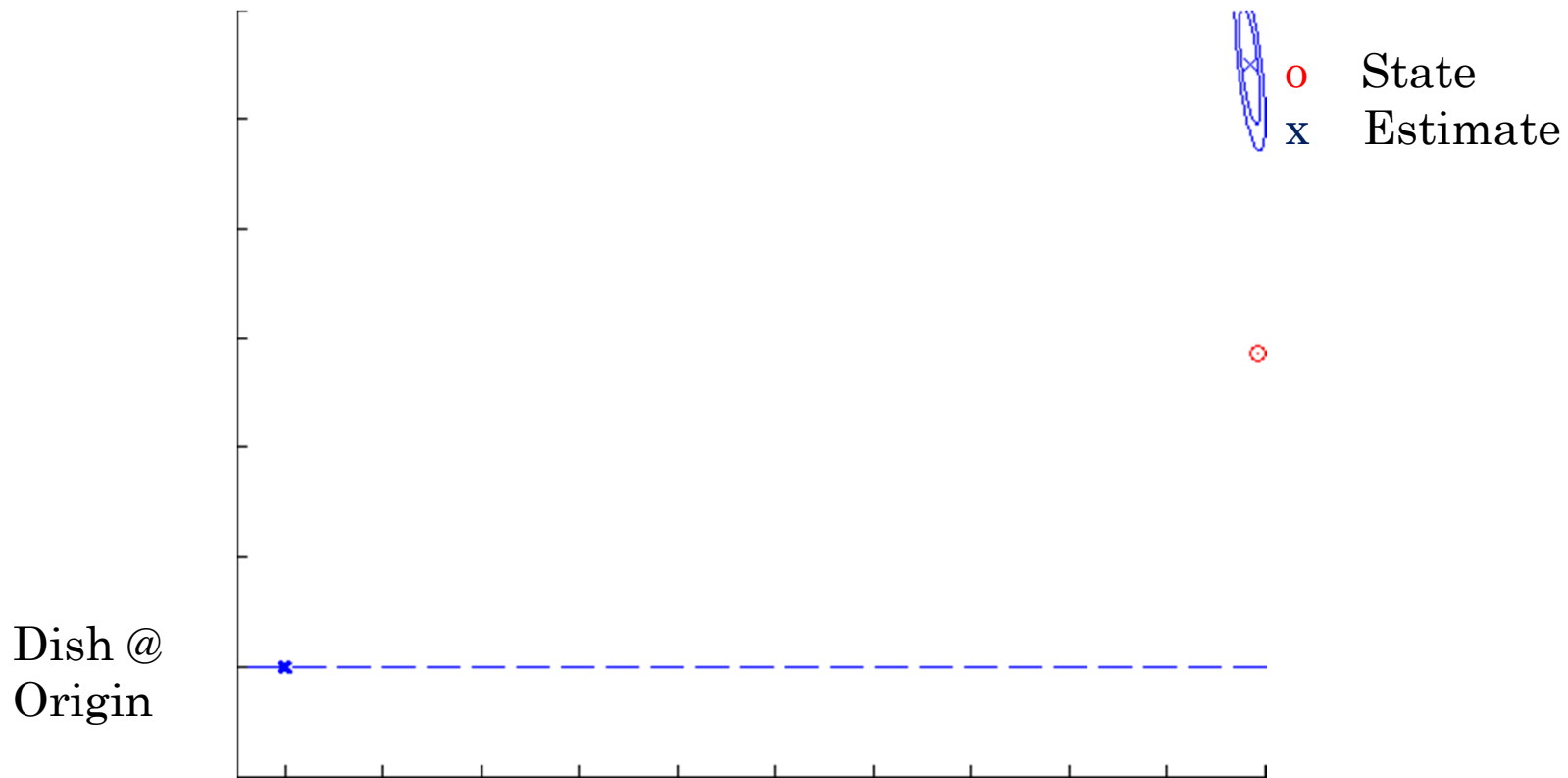


EXTENDED KALMAN FILTER

○ Results

- Noisy noise, big disturbances, incorrect prior

$$\mu_0 = [22 \quad -1.8 \quad 6]^T$$

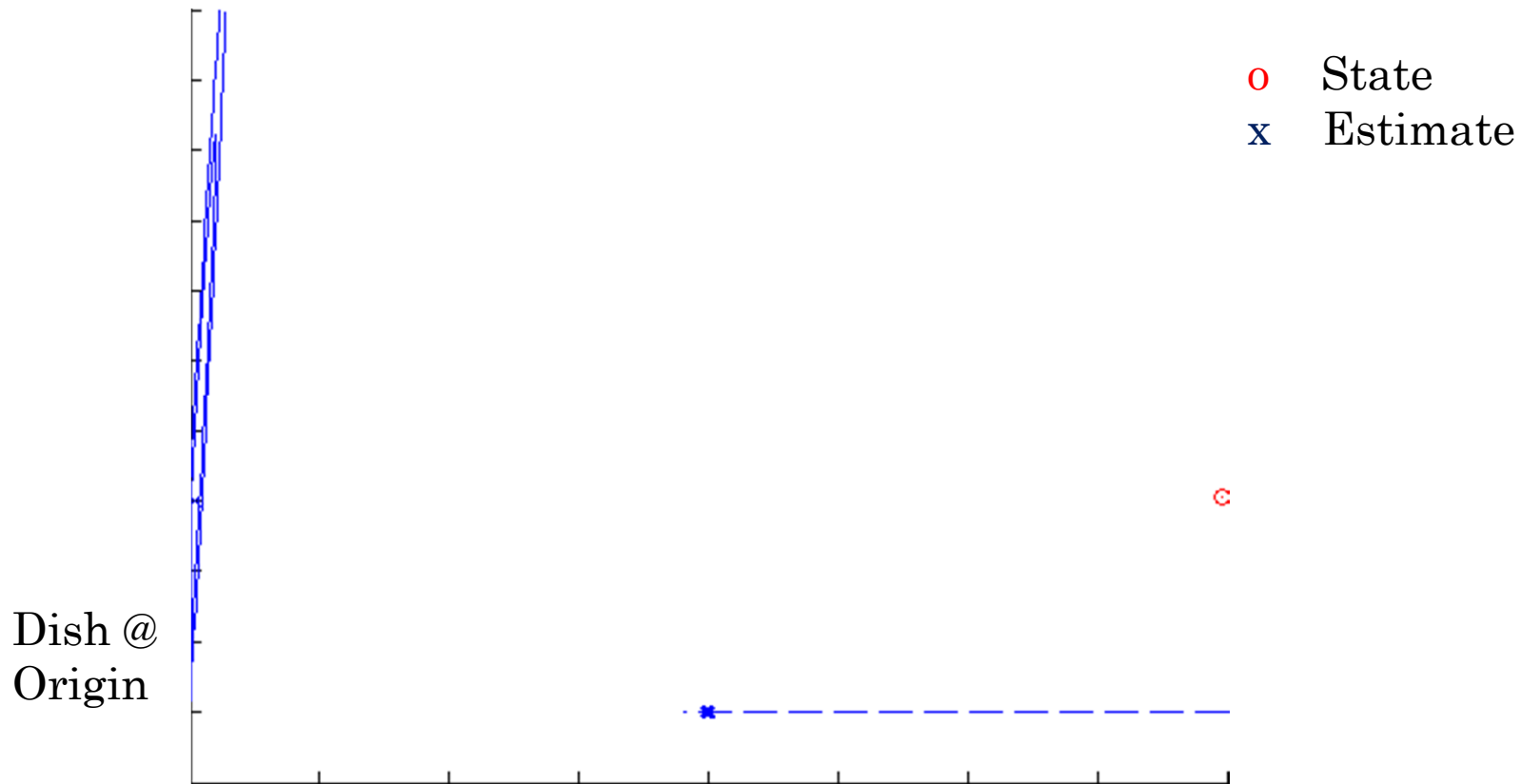


EXTENDED KALMAN FILTER

○ Results

- Symmetrically incorrect prior

$$\mu_0 = [-20 \quad 2 \quad 3]^T$$



EXTENDED KALMAN FILTER

○ Summary

- Direct extension of KF to nonlinear models
- Use Taylor series expansion to find locally linear approximations
- No longer optimal
- Most effective when covariance is low
 - Local linear approximation more likely to be accurate over range of distribution
- Covariance update may diverge

EXTRA SLIDES

KALMAN FILTER

- Reminder on generating multivariate random noise samples

- Define two distributions, the one of interest and the standard normal distribution

$$\delta \sim N(\mu, \Sigma) \qquad \omega \sim N(0, I)$$

- If the covariance is full rank, it can be diagonalized
 - Symmetry implies positive semidefiniteness

$$\begin{aligned}\Sigma &= E\lambda E^T \\ &= E\lambda^{1/2} I \lambda^{1/2} E^T \\ &= H I H^T\end{aligned}$$

- Can now relate the two distributions (linear identity)

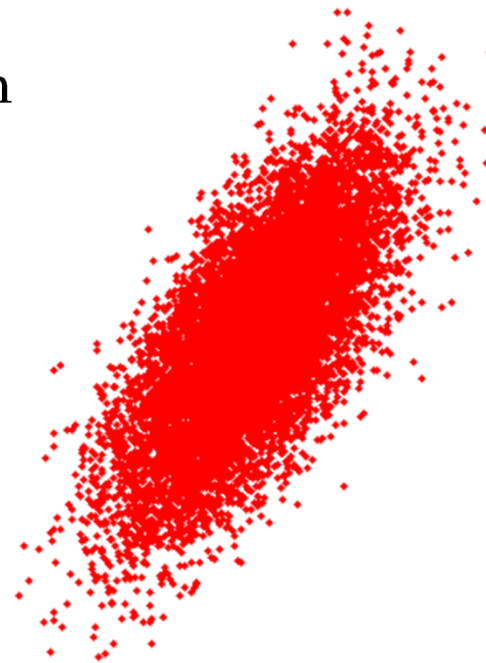
$$\begin{aligned}\delta &\sim N(\mu, H I H^T) \\ \delta &= \mu + H\omega\end{aligned}$$

KALMAN FILTER

- To implement this in Matlab for simulation purposes
 - Define μ, Σ
 - Find eigenvalues, λ , and eigenvectors, E of Σ
 - The noise can then be created with

$$\delta = \mu + E\lambda^{1/2}\text{randn}(n,1)$$

$$\Sigma = \begin{bmatrix} 4 & 4 \\ 4 & 8 \end{bmatrix}$$

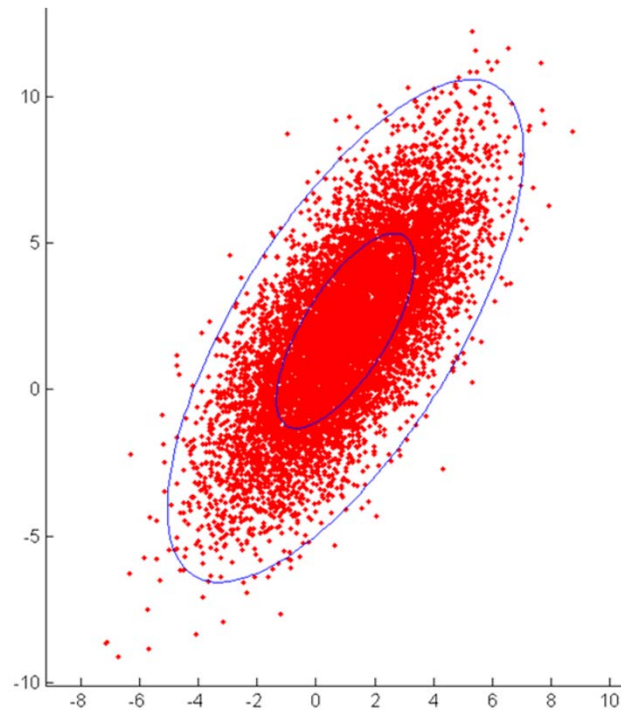


KALMAN FILTER

- Note on confidence ellipses
 - Lines of constant probability
 - Found by setting pdf exponent to a constant
 - Principal axes are eigenvectors of covariance
 - Magnitudes depend on eigenvalues of covariance

$$\mu = \begin{bmatrix} 1 \\ 2 \end{bmatrix}, \quad \Sigma = \begin{bmatrix} 4 & 4 \\ 4 & 8 \end{bmatrix}$$

50%, 99% error ellipses
Not easily computed,
code provided



UNSCENTED KALMAN FILTER

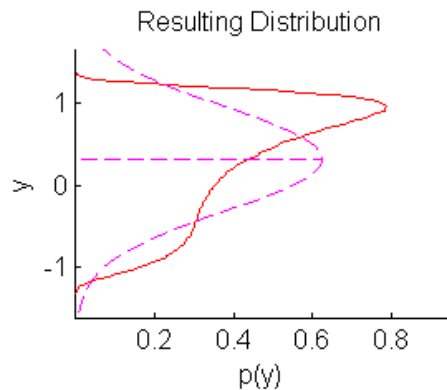
- The EKF used linearization about the predicted/previous state estimate to update the mean and covariance of the current estimate
 - Approximation of a nonlinear transformation of a Gaussian distribution by linear transformation of the mean and covariance
- There are other ways to approximate this transformation
 - Unscented transform leads to better estimates of resulting mean and covariance in some cases
 - Relies on a set of samples known as sigma points or particles, that get transformed directly
 - UKF first published in 1997, still being discussed, extended, solidified.

UNSCENTED KALMAN FILTER

- Key idea: Unscented transform
 - It is more accurate to approximate a distribution using samples than it is to approximate an arbitrary nonlinear function through linearization.
 - Let's first go back to the nonlinear function of a Gaussian and see what the EKF is doing.

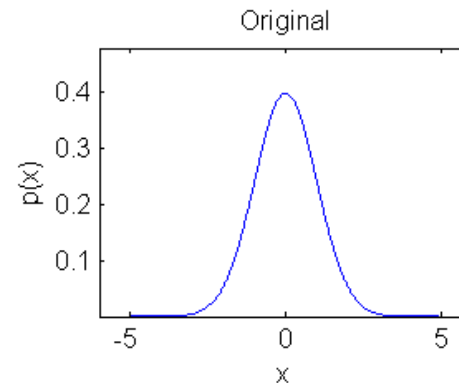
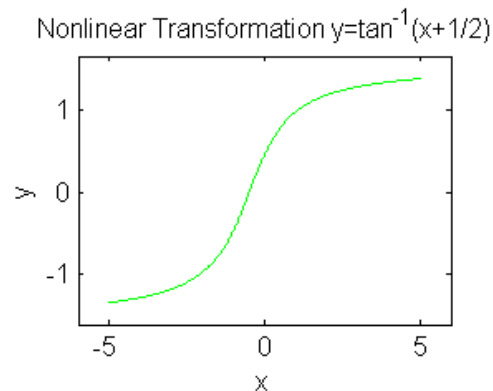
UNSCENTED TRANSFORM

- Effect of nonlinearity on Gaussian distribution
 - Nonlinear transformation



$y \sim N(\mu, \Sigma)$??

$$y = \tan^{-1}(x + 1/2)$$



$$x \sim N(0, 1)$$

UNSCENTED TRANSFORM

- Nonlinear distribution generation

- Take 5,000,000 samples of original Gaussian

$$x^i \sim N(0,1), \quad i = 1, \dots, n$$

- Apply nonlinear transformation to each sample

$$y^i = \tan^{-1}(x^i + 1/2)$$

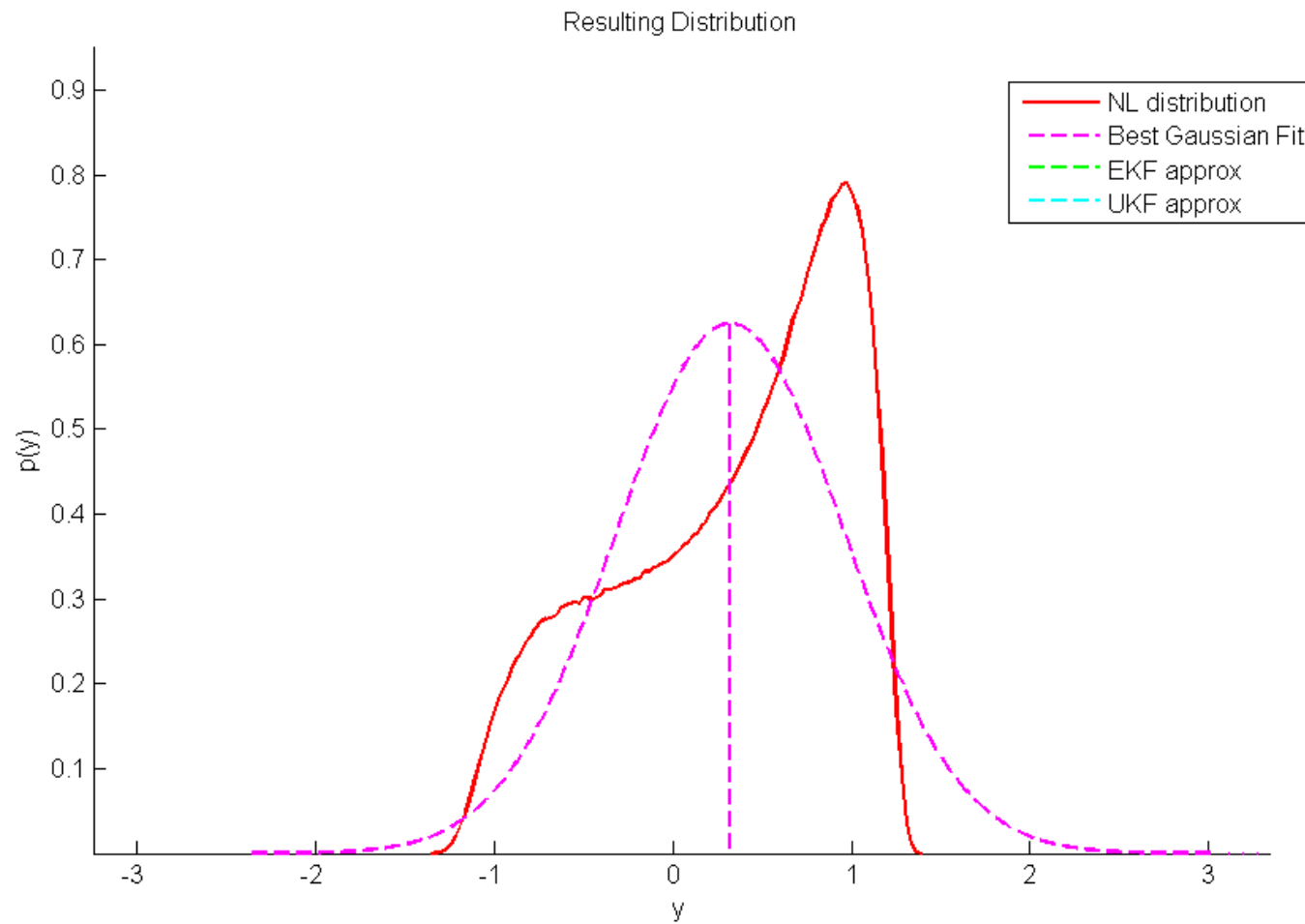
- Create histogram with 100 bins and normalize counts

- Best Gaussian fit generation

- Calculate mean and covariance of 5,000,000 transformed samples

$$\mu_{BG} = \frac{1}{n} \sum_{i=1}^n y^i \quad \Sigma_{BG} = \frac{1}{n-1} \sum_{i=1}^n (y^i - \mu_{BG})(y^i - \mu_{BG})^T$$

UNSCENTED TRANSFORM



UNSCENTED TRANSFORM

- Extended Kalman Filter approximation generation
 - Linearize nonlinear function about mean

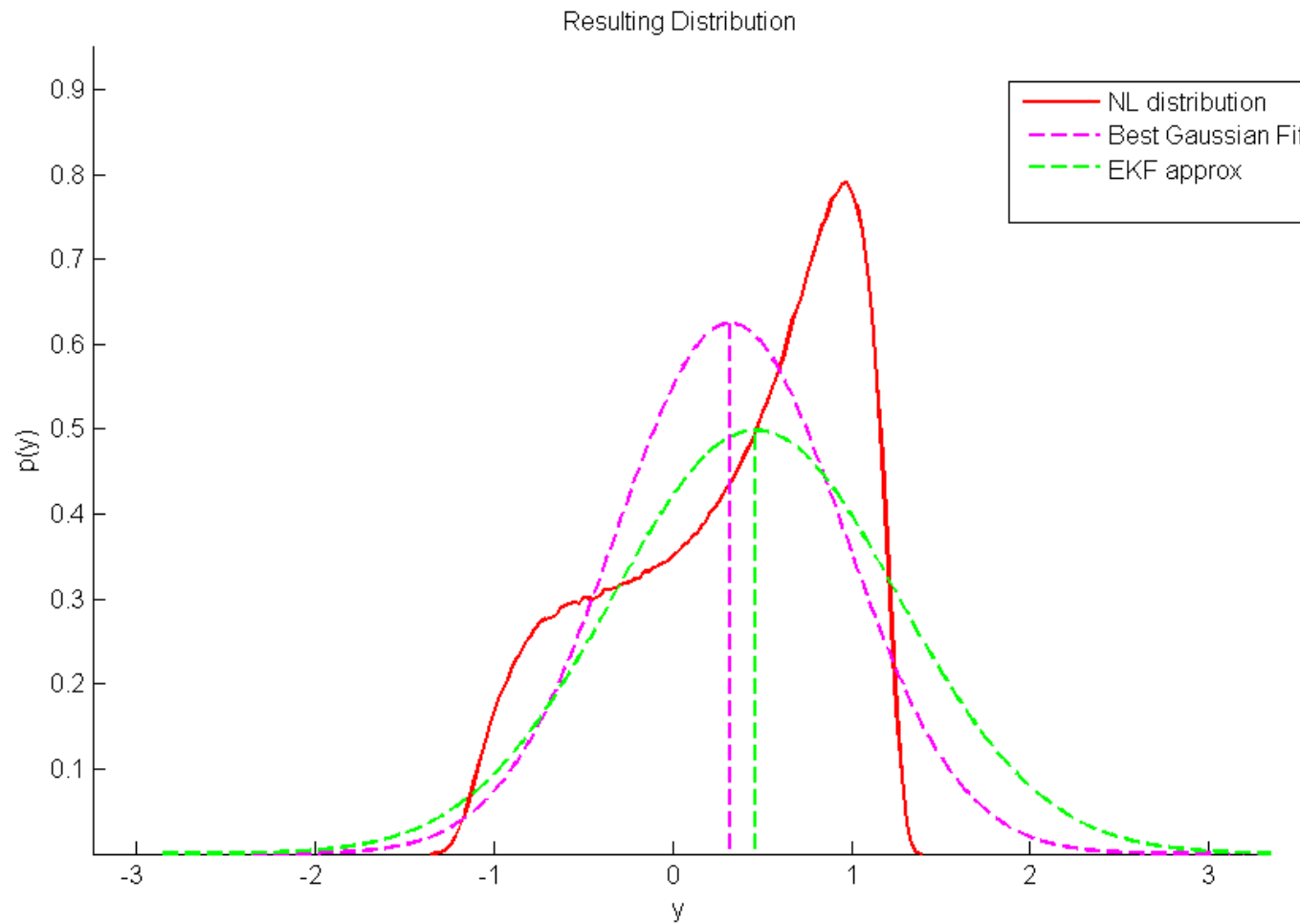
$$G = \left. \frac{\partial}{\partial x} \left(\tan^{-1} (x + 1/2) \right) \right|_{x=\mu}$$
$$= \frac{1}{(\mu + 1/2)^2 + 1}$$

- Propagate mean through nonlinear function and covariance through linearized function

$$\mu_{EKF} = \tan^{-1} (\mu + 1/2)$$

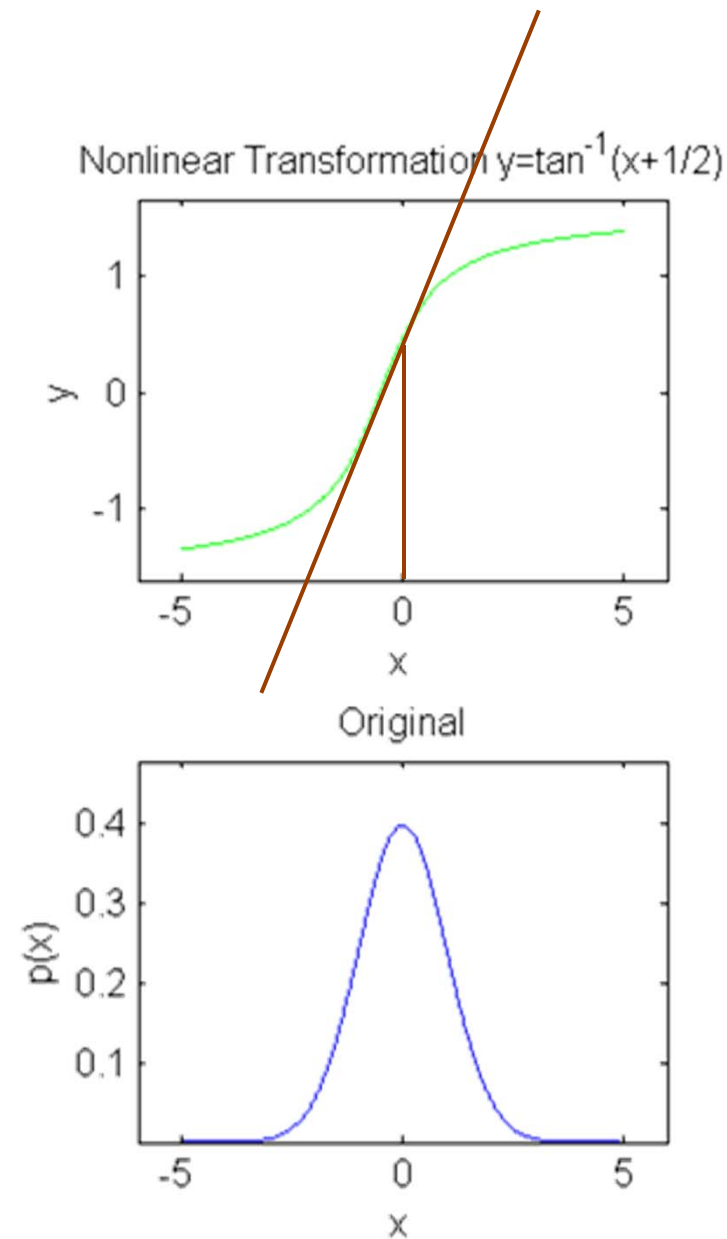
$$\Sigma_{EKF} = G \Sigma G^T$$

UNSCENTED KALMAN FILTER



LINEARIZATION

- Linearization over-predicts mean shift
 - Assumes symmetry of atan
- Covariance over-predicted as well
 - atan has effect of piling up tails at +/- 1.57



UNSCENTED TRANSFORM

- The unscented transform can also be used
 - Linearization is a first order approximation
 - The unscented transform is second order accurate, and can be tuned to reduce fourth order errors
- The transform relies on a set of specially chosen samples known as sigma points
 - $2n+1$ points chosen to capture the transformation of the distribution

UNSCENTED TRANSFORM

- In 1D case, the unscented transform select 3 points

$$\chi^{[0]} = \mu$$

$$\chi^{[1]} = \mu + \sigma$$

$$\chi^{[2]} = \mu - \sigma$$

- And we select weights so that we can recover the original mean and variance

$$\mu = \sum_{i=0}^2 w_m^{[i]} \chi^{[i]} \quad \sigma^2 = \sum_{i=0}^2 w_c^{[i]} (\chi^{[i]} - \mu)^2$$

UNSCENTED TRANSFORM

- We then pass the sigma points through the nonlinear function

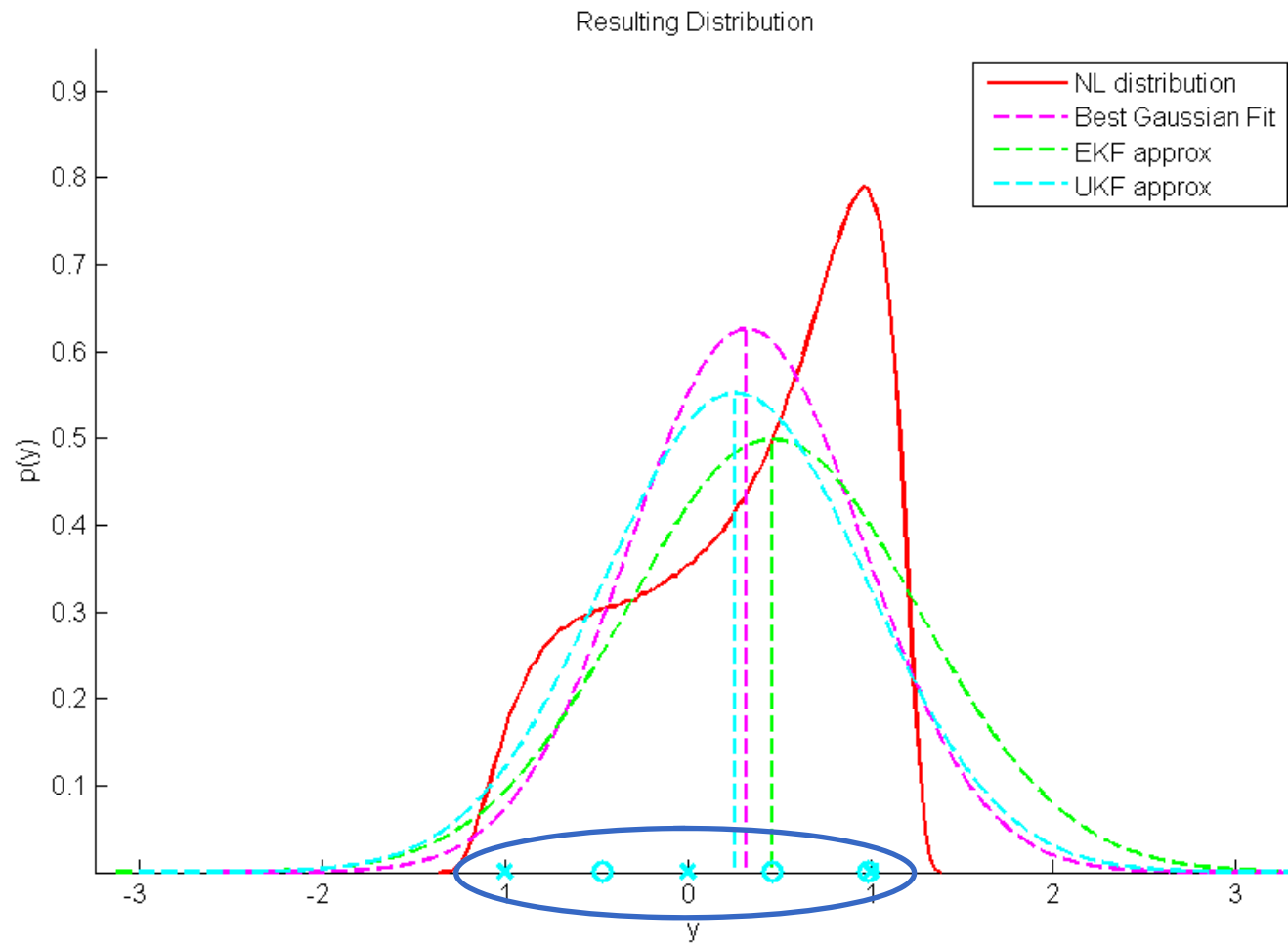
$$\Upsilon^{[i]} = f(\chi^{[i]})$$

- And construct the new mean and variance using the same weights

$$\mu_Y = \sum_{i=0}^2 w_m^{[i]} \Upsilon^{[i]}$$

$$\sigma_Y^2 = \sum_{i=0}^2 w_c^{[i]} (\Upsilon^{[i]} - \mu_Y)^2$$

UNSCENTED TRANSFORM



UNSCENTED TRANSFORM

- In general, the sigma points are chosen as follows

$$\chi^{[0]} = \mu$$

$$\chi^{[i]} = \mu + \left(\sqrt{(n + \lambda) \Sigma} \right)_i, \quad i = 1, \dots, n$$

$$\chi^{[n+i]} = \mu - \left(\sqrt{(n + \lambda) \Sigma} \right)_i, \quad i = 1, \dots, n$$

- Generalized Std. Dev. is square root of covariance
- Here the square root of the covariance matrix is ambiguous, but must satisfy $A = \sqrt{B} \Rightarrow A^T A = B$
 - Can use sqrtm, which returns the unique solution with non-negative eigenvalues,
 - Or use chol, the cholesky decomposition, which returns an upper triangular square root and is very efficient
 - Assumes symmetry

UNSCENTED TRANSFORM

- The parameter λ defines the weights to use for generating the mean and covariance, can be tuned

$$\lambda = \alpha^2 (n + \kappa) - n$$

- α governs the spread of the sigma points about the mean
 - the larger the α the larger the spread of sigma points
 - Usually, $0 \leq \alpha \leq 1$
- κ ensures positive semi-definiteness if $\kappa \geq 0$
 - Can be left at 0 safely (ignored)
 - Also affects the spread of sigma points

UNSCENTED TRANSFORM

- The sigma points are then propagated through the nonlinear function

$$\Upsilon^{[i]} = f(\chi^{[i]})$$

- And a mean and covariance is extracted using special weights for the sigma points

$$\mu_Y = \sum_{i=0}^{2n} w_m^{[i]} \Upsilon^{[i]}$$

$$\Sigma_Y = \sum_{i=0}^{2n} w_c^{[i]} (\Upsilon^{[i]} - \mu_Y)(\Upsilon^{[i]} - \mu_Y)^T$$

UNSCENTED TRANSFORM

- The weights are defined as

$$w_m^{[0]} = \frac{\lambda}{n + \lambda}$$

$$w_c^{[0]} = \frac{\lambda}{n + \lambda} + 1 - \alpha^2 + \beta$$

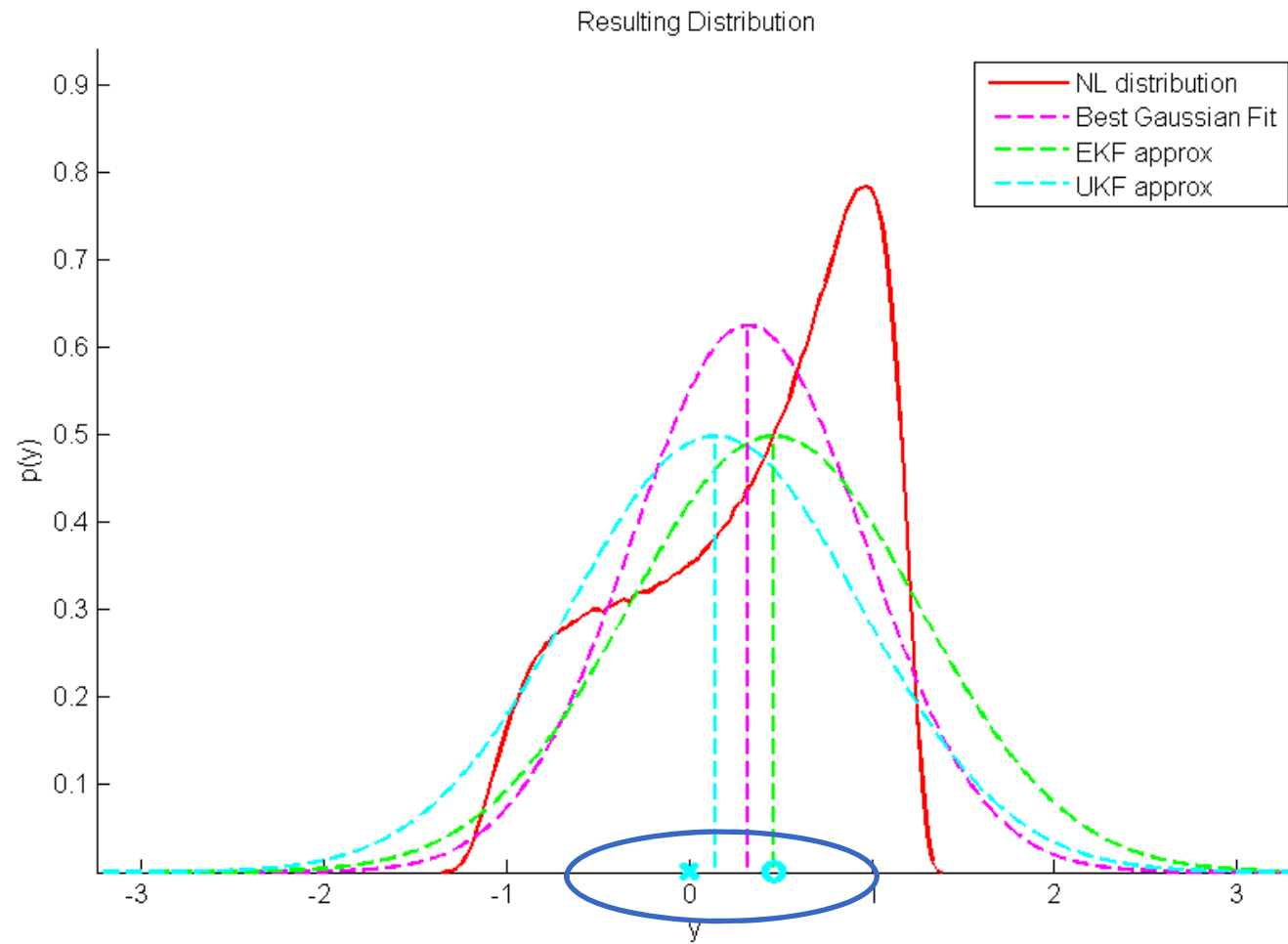
$$w_m^{[i]} = \frac{1}{2(n + \lambda)}, i = 1, \dots, 2n$$

$$w_c^{[i]} = \frac{1}{2(n + \lambda)}, i = 1, \dots, 2n$$

- With another tunable parameter β ,
 - Can be ignored
 - Or set to 2
 - reduces errors in some of the fourth order terms for a Gaussian prior

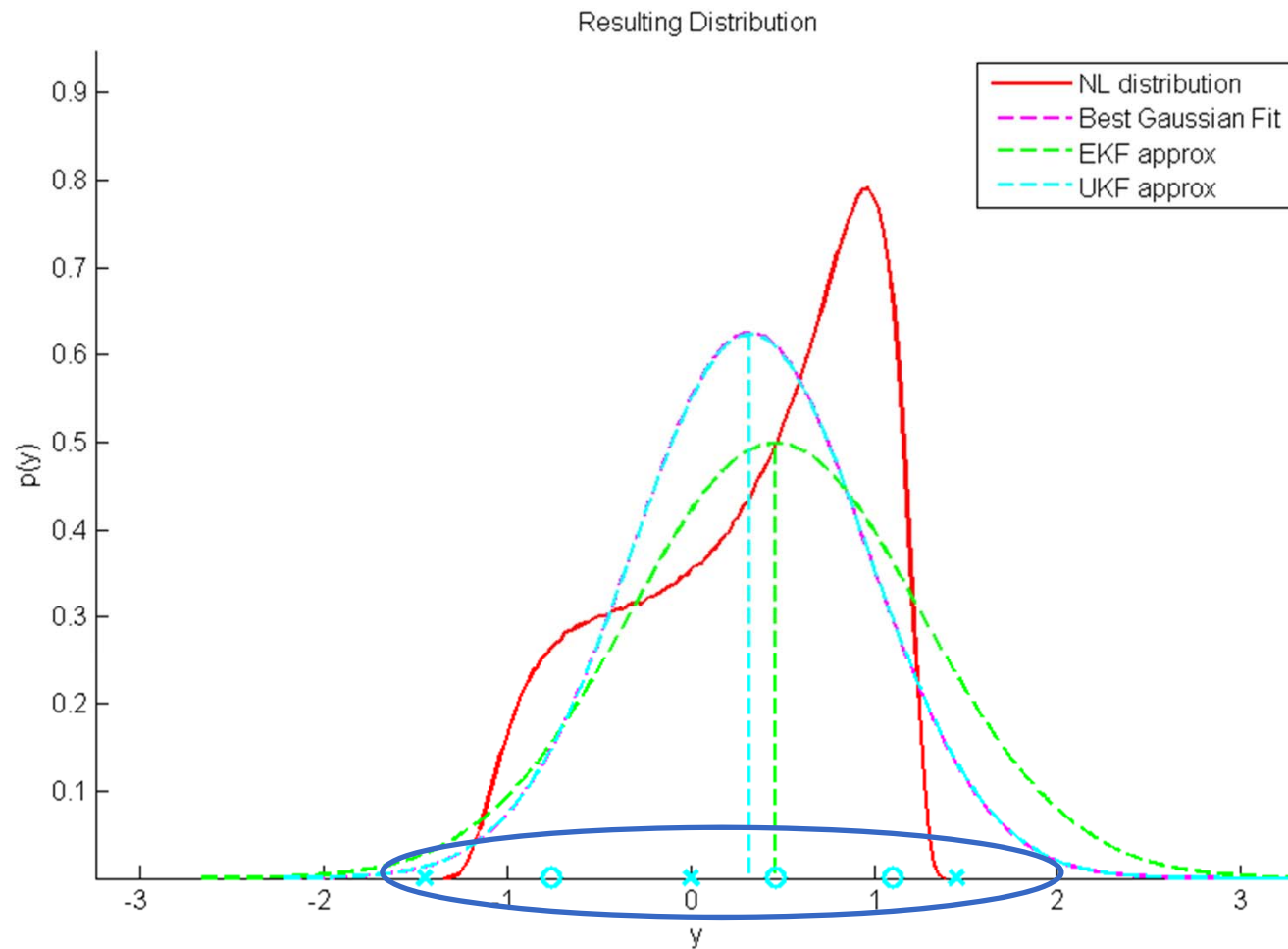
UNSCENTED TRANSFORM

- Select $\alpha = 0.01, \kappa = 0, \beta = 0$



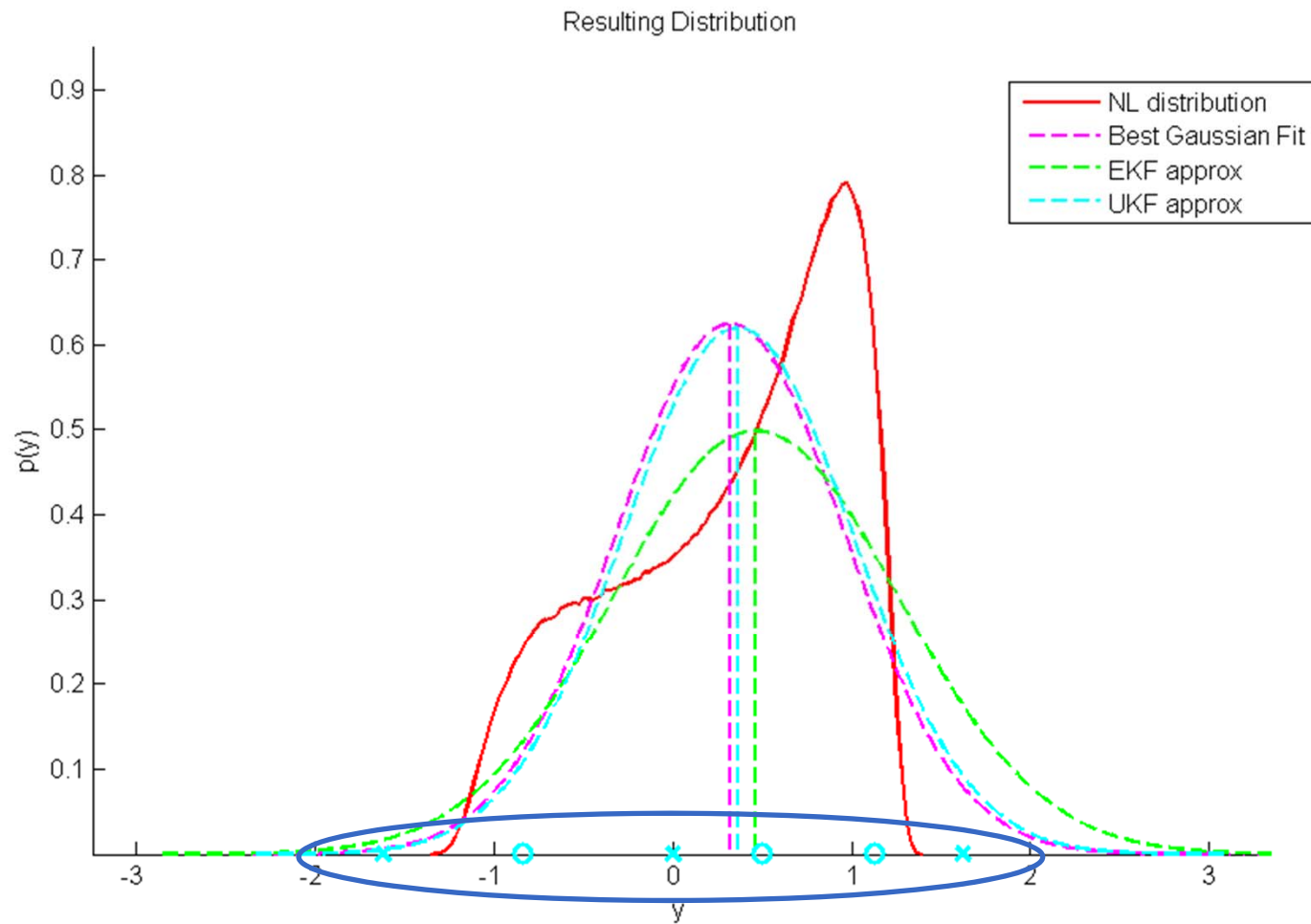
UNSCENTED TRANSFORM

- Select $\alpha = 1.45, \kappa = 0, \beta = 0$



UNSCENTED TRANSFORM

- Select $\alpha = 1.63, \kappa = 0, \beta = 2$

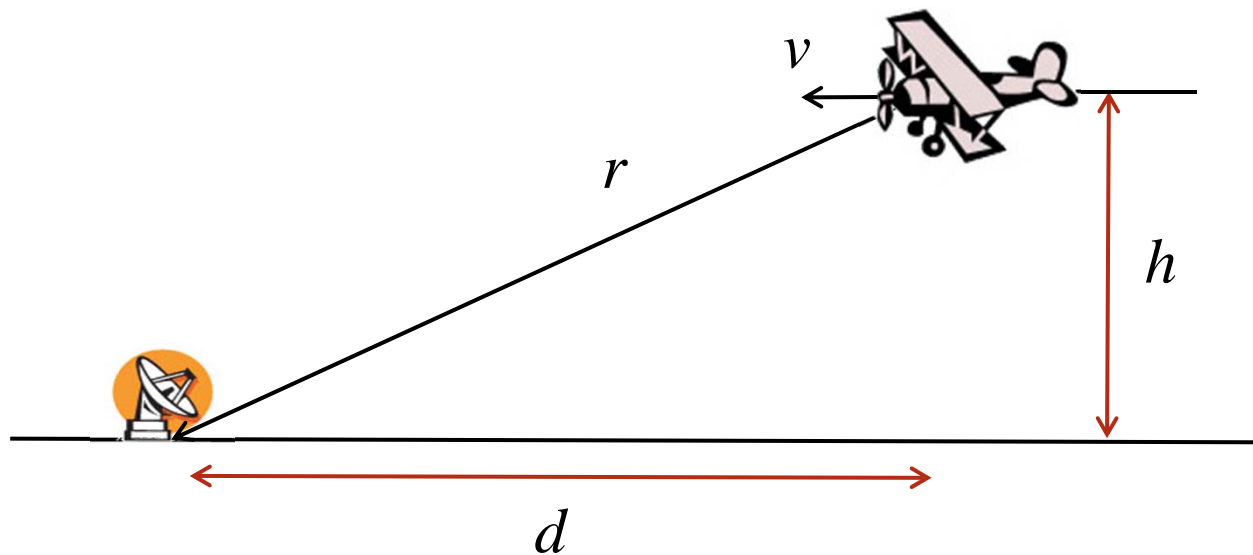


UNSCENTED KALMAN FILTER

- Incorporating this method of distribution transformation into the Bayesian framework is possible
 - There are two nonlinear functions to deal with
 - Two unscented transforms are needed per timestep
 - The measurement model depends on the state we are trying to estimate
 - The state is augmented by the measurement noise states and a joint probability density function is updated

UNSCENTED KALMAN FILTER

- Example repeat
 - Radar measurement of an airplane position while flying at constant altitude and velocity



UNSCENTED KALMAN FILTER

○ Example

- State

$$x = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} d \\ v \\ h \end{bmatrix}$$

Initial

$$x_0 = \begin{bmatrix} 20 \\ -2 \\ 3 \end{bmatrix}$$

- Motion Model

$$x_{1,t} = x_{1,t-1} + x_{2,t-1} dt$$

$$x_{2,t} = x_{2,t-1}$$

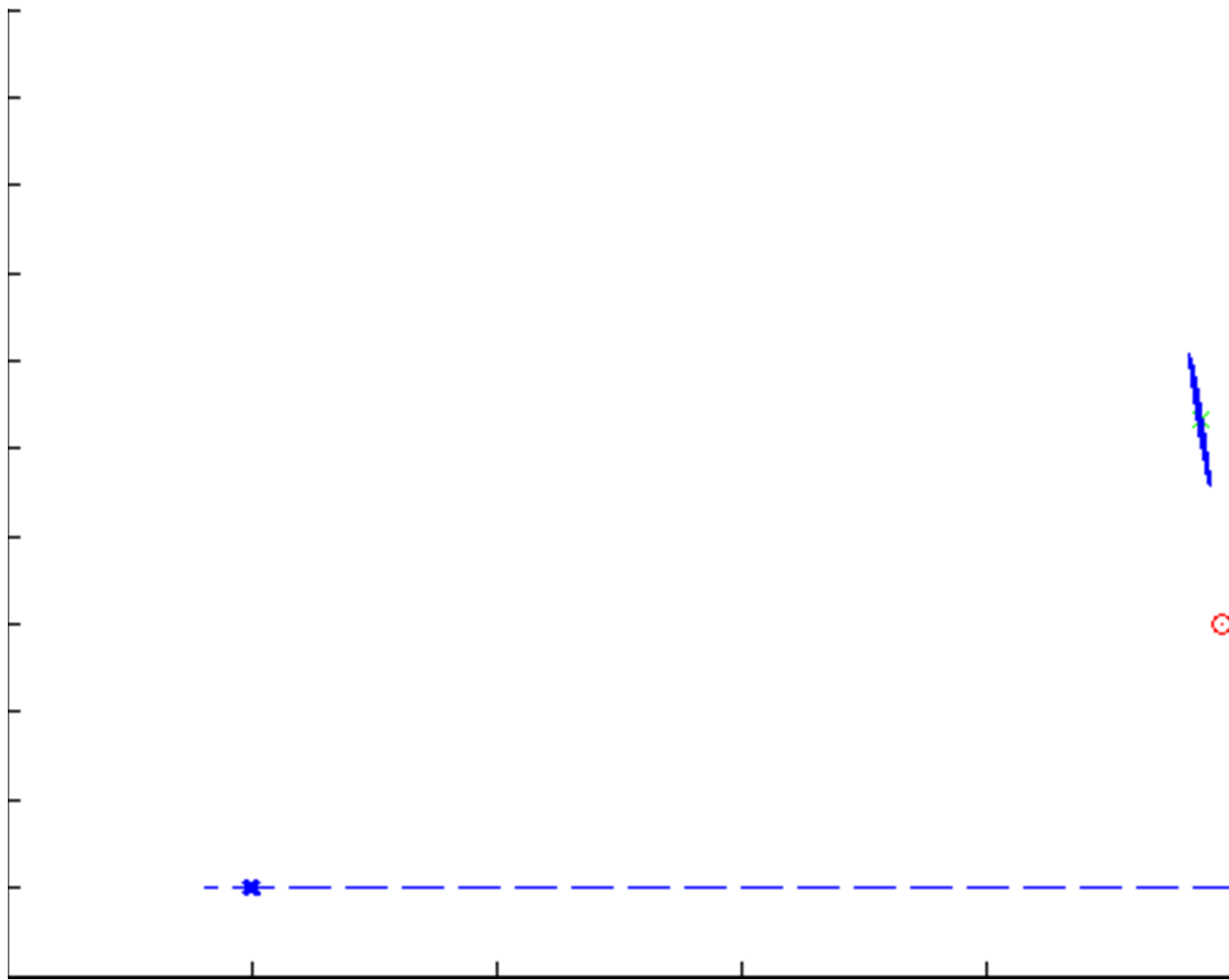
$$x_{3,t} = x_{3,t-1}$$

- Measurement Model

$$y_t = \sqrt{x_{1,t}^2 + x_{3,t}^2} + \delta_t$$

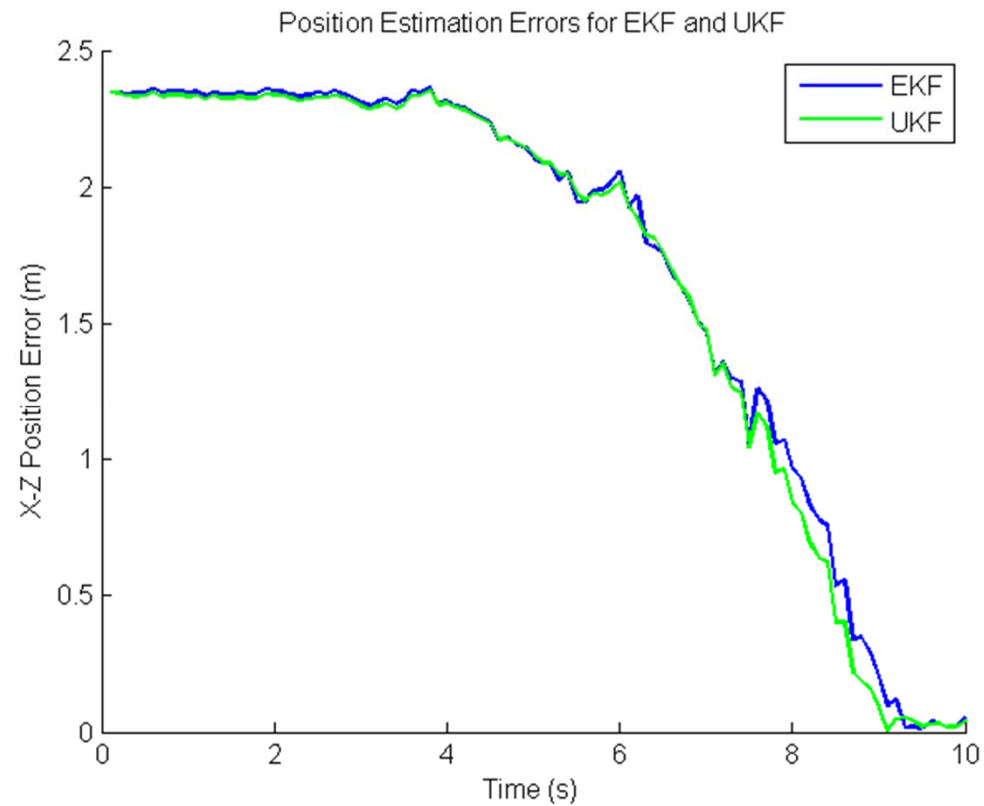
UNSCENTED KALMAN FILTER

- Simulation results- low disturbances, noise



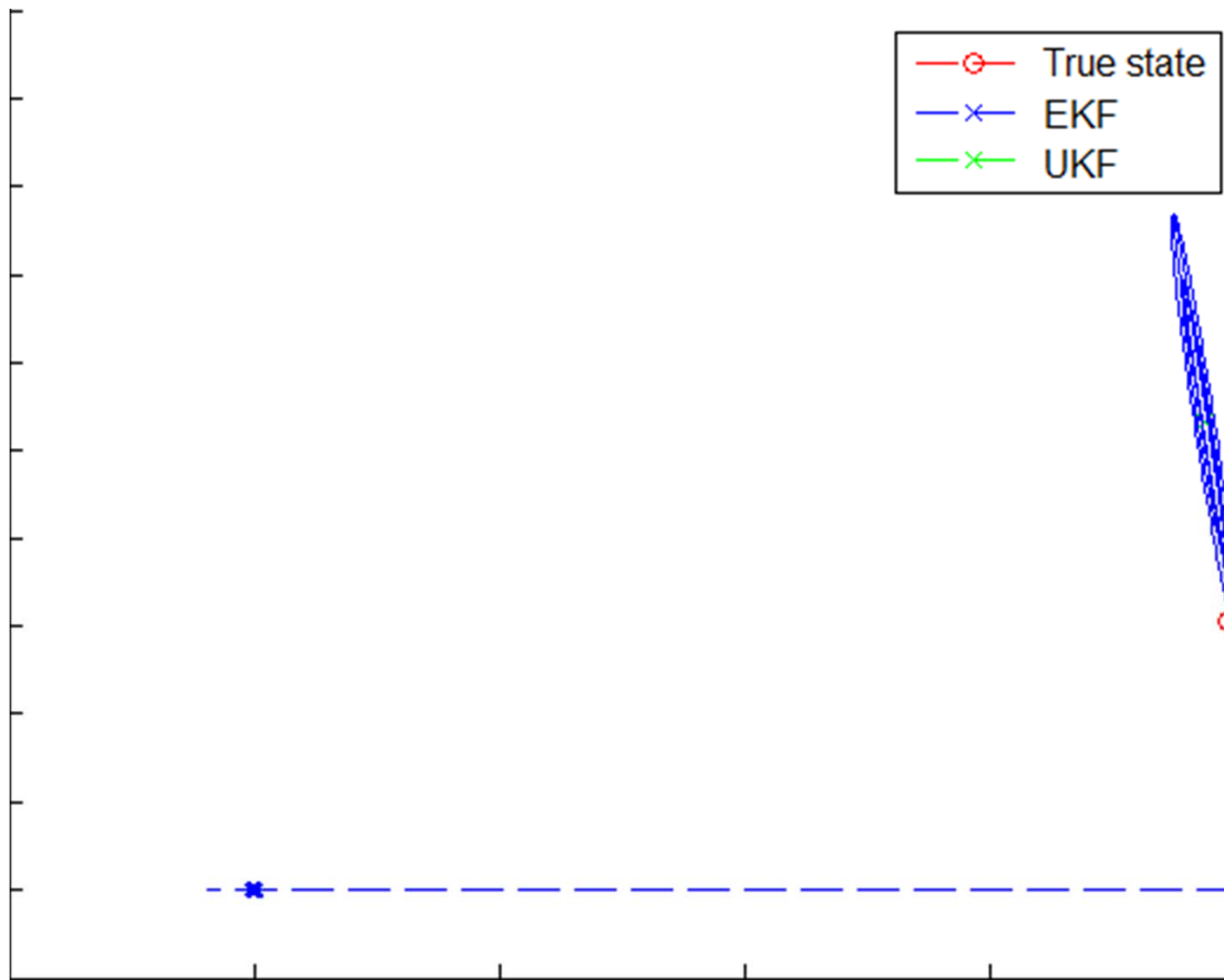
UNSCENTED KALMAN FILTER

- Error plot for position error



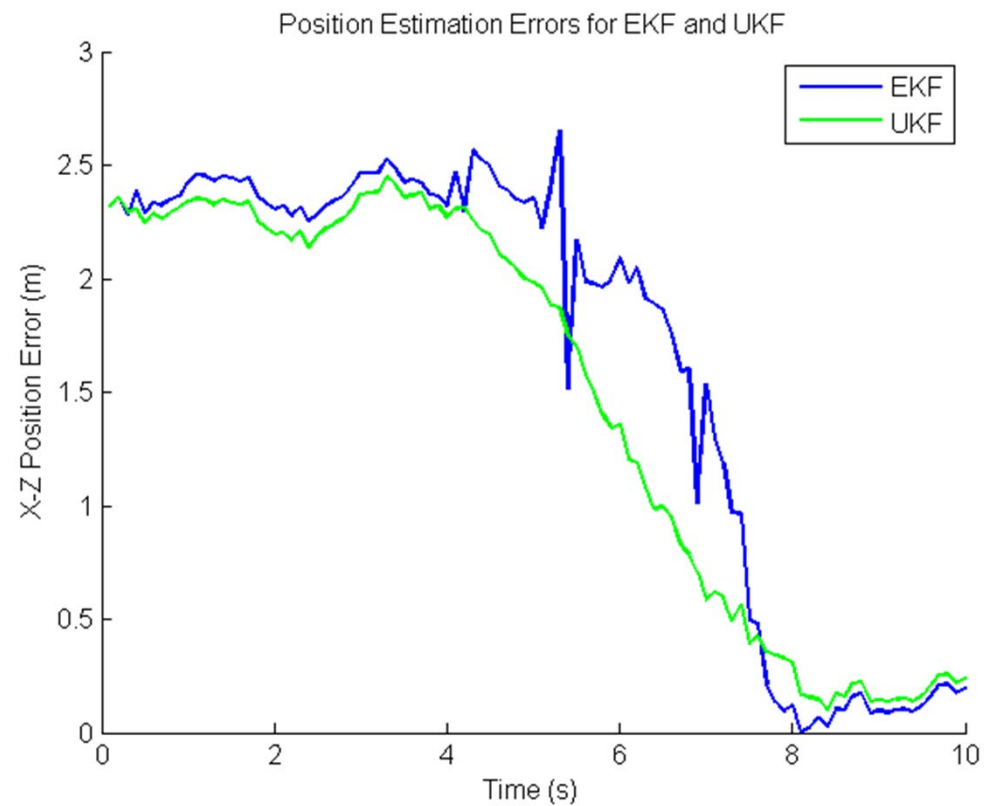
UNSCENTED KALMAN FILTER

- Simulation results, higher disturbances



UNSCENTED KALMAN FILTER

- Error plot for position error



UNSCENTED KALMAN FILTER

- Unscented Kalman Filter Modeling Assumptions

- Prior over the state is Gaussian

$$p(x_0) \sim N(\mu_0, \Sigma_0)$$

- Motion model, nonlinear but still with additive Gaussian disturbances

$$x_t = g(x_{t-1}, u_t) + \varepsilon_t \quad \varepsilon_t \sim N(0, R_t)$$

- Measurement model also nonlinear with additive Gaussian noise

$$y_t = h(x_t) + \delta_t \quad \delta_t \sim N(0, Q_t)$$

- Nonlinearity destroys certainty that beliefs remain Gaussian

UNSCENTED KALMAN FILTER

- Prediction step

- Propagation of belief at $t-1$ through motion model
 - Pick sigma points

$$\chi_{t-1}^{[0]} = \mu_{t-1}$$

$$\chi_{t-1}^{[i]} = \mu_{t-1} + \left(\sqrt{(n + \lambda) \Sigma_{t-1}} \right)_i, \quad i = 1, \dots, n$$

$$\chi_{t-1}^{[n+i]} = \mu_{t-1} - \left(\sqrt{(n + \lambda) \Sigma_{t-1}} \right)_i, \quad i = 1, \dots, n$$

- Propagate through motion model

$$\bar{\chi}_t^{[i]} = g \left(\chi_{t-1}^{[i]}, u_t \right)$$

UNSCENTED KALMAN FILTER

○ Prediction step

- Unscented prediction step
 - Calculate mean and covariance, adding motion covariance to result

$$\bar{\mu}_t = \sum_{i=0}^{2n} w_m^{[i]} \bar{\chi}_t^{[i]}$$

$$\bar{\Sigma}_t = \sum_{i=0}^{2n} w_c^{[i]} \left(\bar{\chi}_t^{[i]} - \bar{\mu}_t \right) \left(\bar{\chi}_t^{[i]} - \bar{\mu}_t \right)^T + R_t$$

UNSCENTED KALMAN FILTER

- Measurement step

- Recall from Bayes filter, we are trying to define

$$bel(x_t) = \eta p(y_t | x_t) \overline{bel}(x_t)$$

- We have the mean and covariance of predicted belief
- We need to propagate this belief through another unscented transform
 - To do this, we need to look at the joint unscented transform

UNSCENTED KALMAN FILTER

Joint transform

- with additive noise in model

$$\delta \sim N(0, Q)$$

$$\begin{bmatrix} x \\ y = h(x) + \delta \end{bmatrix} \sim N\left(\begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}, \begin{bmatrix} \Sigma_x & \Sigma_{xy} \\ \Sigma_{xy} & \Sigma_y \end{bmatrix}\right)$$

- Mean and covariance is found as before
 - Generate sigma points
 - Propagate through model
 - Find mean as before and covariance, cross-covariance as

$$\Sigma_Y = \sum_{i=0}^{2n} w_c^{[i]} (\Upsilon^{[i]} - \mu_Y)(\Upsilon^{[i]} - \mu_Y)^T + Q$$

$$\Sigma_{\chi Y} = \sum_{i=0}^{2n} w_c^{[i]} (\chi^{[i]} - \mu_\chi)(\Upsilon^{[i]} - \mu_Y)^T$$

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- Magic trick (Schur's complement)

- If

$$\begin{bmatrix} x \\ y \end{bmatrix} \sim N \left(\begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}, \begin{bmatrix} \Sigma_x & \Sigma_{xy} \\ \Sigma_{xy} & \Sigma_y \end{bmatrix} \right)$$

- Then

$$p(x|y) = N \left(\mu_x + \Sigma_y \Sigma_{xy}^{-1} (y - \mu_y), \Sigma_x - \Sigma_y \Sigma_{xy}^{-1} \Sigma_y^T \right)$$

- Can in fact derive KF updates using this as well

UNSCENTED KALMAN FILTER

○ Measurement Step

- Form the joint distribution of x_t, y_t given all inputs and all but the latest measurement

$$p(x_t, y_t | y_{1:t-1}, u_{1:t}) = N \left(\begin{bmatrix} \mu_{x_t} \\ \mu_{y_t} \end{bmatrix}, \begin{bmatrix} \Sigma_{x_t} & \Sigma_{x_t y_t} \\ \Sigma_{x_t y_t} & \Sigma_{y_t} \end{bmatrix} \right)$$

- Solve for all components and apply Schur's complement
- But some of these we know already

$$p(x_t | y_{1:t-1}, u_{1:t}) = \overline{bel}_t \quad \longrightarrow \quad \begin{aligned} \mu_{x_t} &= \overline{\mu}_t \\ \Sigma_{x_t} &= \overline{\Sigma}_t \end{aligned}$$

UNSCENTED KALMAN FILTER

- Measurement step

- The rest we can approximate with the unscented transform
- Generate new sigma points from the predicted belief

$$\bar{\chi}_t^{[0]} = \bar{\mu}_t$$

$$\bar{\chi}_t^{[i]} = \bar{\mu}_t + \left(\sqrt{(n + \lambda) \bar{\Sigma}_t} \right)_i, \quad i = 1, \dots, n$$

$$\bar{\chi}_t^{[n+i]} = \bar{\mu}_t - \left(\sqrt{(n + \lambda) \bar{\Sigma}_t} \right)_i, \quad i = 1, \dots, n$$

- Propagate through measurement model

$$\Upsilon_t^{[i]} = h\left(\bar{\chi}_t^{[i]}\right)$$

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○ Measurement step

- Then the measurement terms and cross terms can be approximated as

$$\mu_{y_t} \approx \sum_{i=0}^{2n} w_m^{[i]} \Upsilon_t^{[i]}$$

$$\Sigma_{y_t} \approx \sum_{i=0}^{2n} w_c^{[i]} \left(\Upsilon^{[i]} - \mu_{y_t} \right) \left(\Upsilon^{[i]} - \mu_{y_t} \right)^T + Q_t$$

$$\Sigma_{x_t y_t} \approx \sum_{i=0}^{2n} w_c^{[i]} \left(\bar{\chi}_t^{[i]} - \mu_{x_t} \right) \left(\Upsilon_t^{[i]} - \mu_{y_t} \right)^T$$

UNSCENTED KALMAN FILTER

○ Measurement Step

- Finally, applying Schur's complement

$$p(x | y) = N\left(\mu_x + \Sigma_y \Sigma_{xy}^{-1} (y - \mu_y), \Sigma_x - \Sigma_y \Sigma_{xy}^{-1} \Sigma_y^T\right)$$

- To the above joint distribution

$$p(x_t | y_{1:t}, u_{1:t}) = \text{bel}(x_t) = N(\mu_t, \Sigma_t)$$

- And therefore,

$$\mu_t = \bar{\mu}_t + \Sigma_{y_t} \Sigma_{x_t y_t}^{-1} (y_t - \mu_{y_t})$$

$$\Sigma_t = \bar{\Sigma}_t - \Sigma_{y_t} \Sigma_{x_t y_t}^{-1} \Sigma_{y_t}^T$$

UNSCENTED KALMAN FILTER

○ Summary

• Prediction Step

$$\chi_{t-1}^{[0]} = \mu_{t-1}$$

$$\chi_{t-1}^{[i]} = \mu_{t-1} + \left(\sqrt{(n + \lambda) \Sigma_{t-1}} \right)_i, \quad i = 1, \dots, n$$

$$\chi_{t-1}^{[n+i]} = \mu_{t-1} - \left(\sqrt{(n + \lambda) \Sigma_{t-1}} \right)_i, \quad i = 1, \dots, n$$

Select sigma points

$$\bar{\chi}_t^{[i]} = g \left(\chi_{t-1}^{[i]}, u_t \right)$$

Apply motion model

$$\bar{\mu}_t = \sum_{i=0}^{2n} w_m^{[i]} \bar{\chi}_t^{[i]}$$

$$\bar{\Sigma}_t = \sum_{i=0}^{2n} w_c^{[i]} \left(\bar{\chi}_t^{[i]} - \bar{\mu}_t \right) \left(\bar{\chi}_t^{[i]} - \bar{\mu}_t \right)^T + R_t$$

Extract mean and covariance

UNSCENTED KALMAN FILTER

Summary

- Measurement step

$$\bar{\chi}_t^{[0]} = \bar{\mu}_t$$

$$\bar{\chi}_t^{[i]} = \bar{\mu}_t + \left(\sqrt{(n + \lambda) \bar{\Sigma}_t} \right)_i, \quad i = 1, \dots, n$$

$$\bar{\chi}_t^{[n+i]} = \bar{\mu}_t - \left(\sqrt{(n + \lambda) \bar{\Sigma}_t} \right)_i, \quad i = 1, \dots, n$$

$$\Upsilon_t^{[i]} = h(\bar{\chi}_t^{[i]})$$

$$\mu_{y_t} \approx \sum_{i=0}^{2n} w_m^{[i]} \Upsilon_t^{[i]}$$

$$\Sigma_{y_t} \approx \sum_{i=0}^{2n} w_c^{[i]} (\Upsilon_t^{[i]} - \mu_{y_t})(\Upsilon_t^{[i]} - \mu_{y_t})^T + Q_t$$

Select sigma points

Apply measurement model

Extract mean and covariance

UNSCENTED KALMAN FILTER

○ Summary

- Measurement Step

$$\Sigma_{x_t y_t} \approx \sum_{i=0}^{2n} w_c^{[i]} (\bar{\chi}_t^{[i]} - \mu_{x_t}) (\Upsilon_t^{[i]} - \mu_{y_t})^T$$

Extract cross-covariance

$$\mu_t = \bar{\mu}_t + \Sigma_{y_t} \Sigma_{x_t y_t}^{-1} (y_t - \mu_{y_t})$$

$$\Sigma_t = \bar{\Sigma}_t - \Sigma_{y_t} \Sigma_{x_t y_t}^{-1} \Sigma_{y_t}^T$$

Update belief using Schur's complement

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○ Summary

- Similar computation time to EKF
 - longer due to square root and inverse
- Potentially capable of reducing errors in propagation of beliefs through nonlinear functions
- Tuning effects unclear, can lead to strange results
- Benefit minimal when nonlinearities are modest, or uncertainty is low